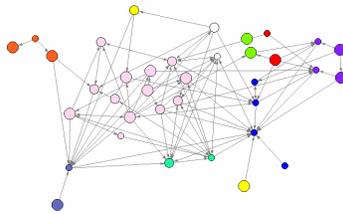


Exponential random graph (p^*) models for social networks

NetSci06



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With acknowledgements to
Pip Pattison, Tom Snijders, Stanley Wasserman, Mark Handcock

Overview

- A. Introduction
 - Why statistical models of social networks?
 - The conceptual basis for exponential random graph models
- B. Bernoulli, Dyadic independence and Markov models
 - Simulation and estimation
- C. Goodness of fit
- D. Degeneracy
- E. New specifications and Goodness of fit
 - Degree sequences
 - Higher order triangulation
- F. Social selection models: incorporating attributes

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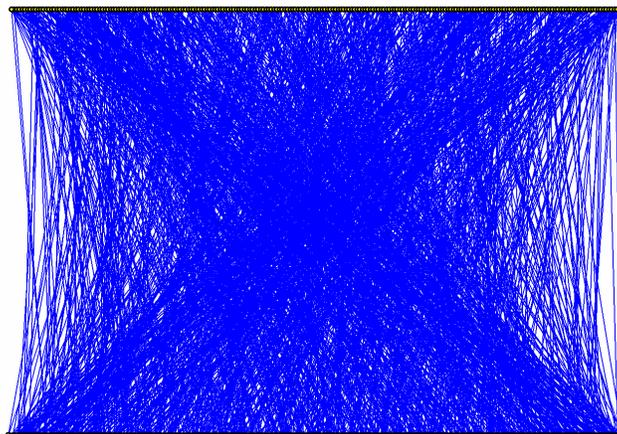
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Why statistical models for social networks?

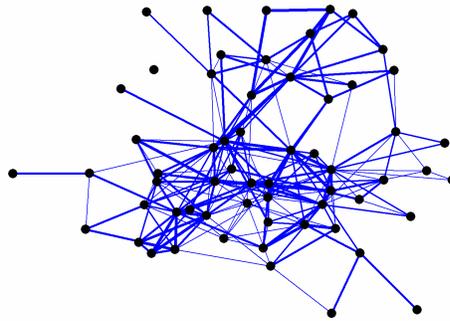


US companies/directors 1996

Why statistical models for social networks?

There are many descriptive techniques that may provide insight into an observed social network:

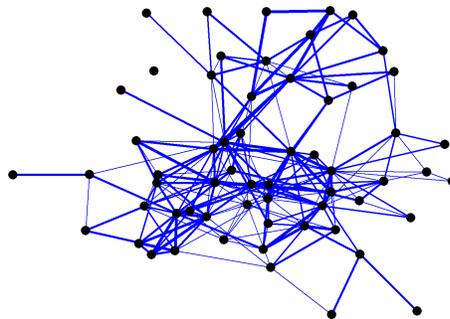
Density, Centrality, Cohesive subsets, etc



Frequency of work interactions

Why statistical models for social networks?

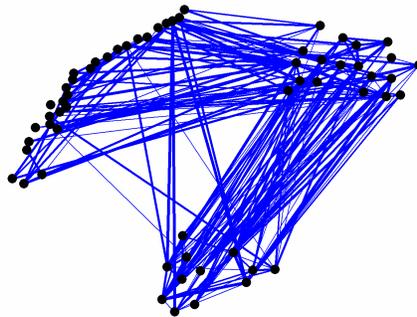
But models may provide some clarity of major effects



Frequency of work interactions

Why statistical models for social networks?

But models may provide some clarity of major effects



Frequency of work interactions

Why statistical models for social networks?

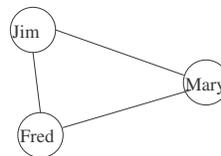
Competing explanations for structural effects.

For instance, the appearance of triangles can occur through different social processes:

1. a *balance-type* effect (triadic closure)
2. an *homophily* effect (people with similar attributes attract).

Only a model including both effects can help decide whether one or both are important.

Friendly Fred introduces Mary to Jim – a balance- or transitivity-type effect.



The outcome is a triangle.

Why statistical models for social networks?

Competing explanations for structural effects.

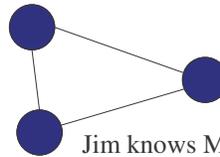
For instance, the appearance of triangles can occur through different social processes:

1. a *balance-type* effect (triadic closure)
2. an *homophily* effect (people with similar attributes attract).

Only a model including both effects can help decide whether one or both are important.

Fred knows Mary
because they support
the same football team.

Fred knows Jim
because they support
the same football team.



Jim knows Mary
because they support
the same football team.

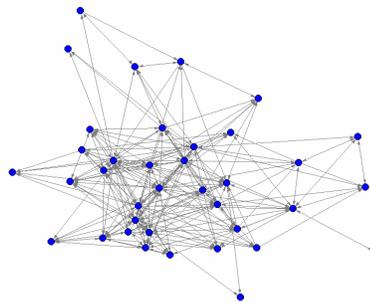
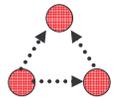
The outcome
is a triangle.

Why statistical models for social networks?

Models may help traverse the micro-macro divide:

What is the range of global outcomes likely from given local network processes?

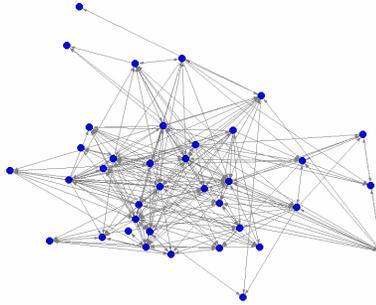
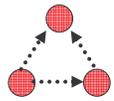
Tendency to transitivity



Why statistical models for social networks?

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*What is the range of **global** outcomes likely from given **local** network processes?*

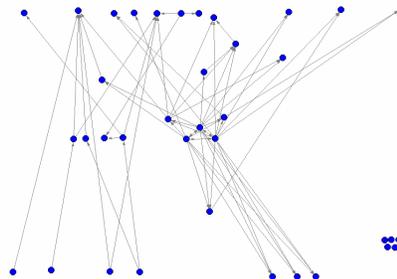
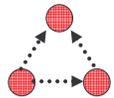
Tendency to transitivity



Why statistical models for social networks?

Models may help traverse the micro-macro divide:
*What is the range of **global** outcomes likely from given **local** network processes?*

Tendency to transitivity



*For an observed network, which **local** network processes may be implicated in its emergence?*

Why statistical models for social networks?

Models may help traverse the micro-macro divide:
*What is the range of **global** outcomes likely
from given **local** network processes?*

And so models (especially statistical models) may help us understand the range of possible outcomes for *processes on networks*

CS

What counts as a “good” statistical model for an observed social network?

1. Models must be **estimable** from data.
2. Model parameters should imply model statistics **consistent** with those of the observed graph.
3. A good model will imply graphs with **other features** that are consistent with the observed graph.
 - *path lengths (geodesic distribution)*
 - *clustering (triangle formation)*
 - *degree distribution*
 - *denser regions (cohesive subsets of nodes)*

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What do we know about processes underlying network tie formation?

Exogenous effects:

Homophily, shared affiliations, spatial propinquity all matter (e.g., McPherson, Smith-Lovin & Cook, 2001)

More generally: network tie formation is often set within *foci* (Feld, 1981) or *settings* (White, 1995; Pattison & Robins, 2002)

Endogenous network effects:

Clustering: tie formation is often more likely when actors have network partners in common (e.g., Cartwright & Harary, 1956; Granovetter, 1973; and many others)

More generally: “a social tie ... is subject to, and known to be subject to, the hegemonic pressures of others engaged in the social construction of that network.” (White, 1998)

Interactions between exogenous and endogenous effects:

General social selection (Robins, Elliott & Pattison, 2001)

Modeling endogenous network processes

Guiding principles:

Network ties are the outcome of (unobserved) social processes that tend to be local and interactive

There are both regularities and irregularities in these local interactive processes

We hence construct statistical models in which:

local interactivity is permitted and assumptions about form of “local interactions” are explicit

regularities are represented by model parameters and estimated from data
consequences of local regularities for global network properties can be understood *and can also provide an exacting approach to model evaluation*

Local interactivity

We model *tie variables*: $\mathbf{X} = [X_{ij}]$ $X_{ij} = 1$ if i has a tie to j
0 otherwise

realization of \mathbf{X} is denoted by $\mathbf{x} = [x_{ij}]$

We also incorporate node-level exogenous *attribute variables*: $\mathbf{Y} = [Y_i]$

Two modeling steps:

Local interactions: define two network tie variables to be *neighbors* if they are conditionally dependent, given the values of all other tie variables

But: what are appropriate *neighbor* assumptions?
Dependency assumptions

**Network dependencies:
which tie variables are neighbors?
Simple dyadic dependence hypotheses**

- (1) There are no neighborhood relations *Bernoulli random graphs*
All edges are independent



- Two tie variables are *neighbors*:
(2) if they are in the same dyad
Directed graphs

dyad independence



**Network dependencies:
which tie variables are neighbors?
More realistic dependence hypotheses**

Two tie variables are *neighbors* if:

- (3) they share an actor



Markov model
(Frank & Strauss, 1986)

- (4) they share connections
with two existing ties
(completing a social circuit)



realization-dependent model
(Pattison & Robins, 2002;
Snijders, Pattison, Robins &
Handcock, 2006)

There are other possibilities, but these two get us a long way

Models for interactive systems of variables (Besag, 1974)

A *neighborhood* is a set of mutually neighboring variables and corresponds to a potential *network configuration*:

e.g. $\{X_{12}, X_{13}, X_{23}\}$ corresponds to



Hammersley-Clifford theorem (Besag, 1974):

A model for \mathbf{X} has a form determined by its neighborhoods

This general approach leads to *exponential random graph* or *p* models*
(Frank & Strauss 1986; extended by Wasserman, Pattison & Robins)

Exponential random graph models

$$P(\mathbf{X} = \mathbf{x}) = (1/c) \exp\{\sum_Q \gamma_Q z_Q(\mathbf{x})\}$$

normalizing quantity

parameter

network statistic

the summation is over all neighbourhoods Q

$z_Q(\mathbf{x}) = \prod_{x_{ij} \in Q} x_{ij}$ signifies whether
all ties in Q are observed in \mathbf{x}

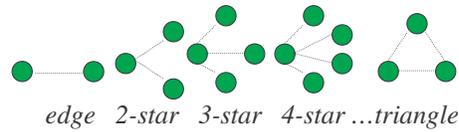
$$c = \sum_{\mathbf{x}} \exp\{\sum_Q \gamma_Q z_Q(\mathbf{x})\}$$

Neighborhoods depend on dependence assumptions

Assumptions: two ties are neighbors: Configurations for neighborhoods

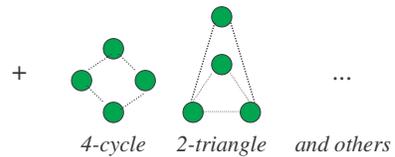
if they share an actor

Markov



if they complete a 4-cycle

realization-dependent



Homogenous models

If we assume that *isomorphic neighbourhoods have equal parameters*, then:

There is one parameter for *each class* of network configurations

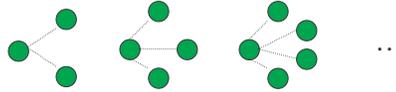
The corresponding statistic is the *number* of configurations in \mathbf{x}

E.g. for a Markov model:

Configurations					...	
Parameters	θ	σ_2	σ_3	σ_4	...	τ
Statistics	$L(\mathbf{x})$	$S_2(\mathbf{x})$	$S_3(\mathbf{x})$	$S_4(\mathbf{x})$...	$T(\mathbf{x})$

Related model parameters

Star configurations



Parameters

σ_2 σ_3 σ_4 ...

If we assume that $\sigma_k = -\sigma_{k-1}/\lambda$, for $k > 1$ and $\lambda \geq 1$ a (fixed) constant
alternating k-star hypothesis

Then we obtain a single *star* parameter (σ_2) with statistic:

$$S^{[\lambda]}(\mathbf{x}) = \sum_k (-1)^k S_k(\mathbf{x}) / \lambda^{k-2} \quad \text{alternating k-star statistic}$$

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**Dependence assumptions:
1. Bernoulli (independent edges)**

Possible edges are independent of one another.
Configurations in this model relate to single possible edges (x_{ij}).



Hammersley-Clifford theorem:

$$\Pr(\mathbf{X} = \mathbf{x}) = \frac{1}{\kappa} \exp\left(\sum \lambda_{ij} x_{ij}\right)$$

**Dependence assumptions:
1. Bernoulli (independent edges)**

One parameter (λ_{ij}) for each possible edge – simply too many.

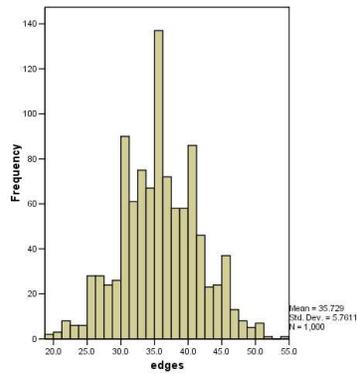
Homogeneity assumption: $\lambda_{ij} = \theta$ for all i, j

Assumes that the edge effect is the same across the entire network.

$$\Pr(\mathbf{X} = \mathbf{x}) = \frac{1}{\kappa} \exp\left(\sum \lambda_{ij} x_{ij}\right) = \frac{1}{\kappa} \exp\left(\theta \sum x_{ij}\right) = \frac{1}{\kappa} \exp(\theta L)$$

where L is the number of edges in the observed network

**Bernoulli graph distribution:
distribution of no. of edges**



**Dependence assumptions:
2. Dyadic independence (directed graphs)**

Dyads are independent of one another.

Configurations in this model relate to arcs (x_{ij}) and reciprocated arcs ($x_{ij} x_{ji}$)



$$\Pr(\mathbf{X} = \mathbf{x}) = \frac{1}{\kappa} \exp(\theta L + \rho M)$$

where L is the number of arcs and M the number of reciprocated arcs in the observed network.

More relaxed homogeneity assumptions: Two-blocks

a. Homogeneous Bernoulli:

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp(\theta L(\mathbf{x}))$$

b. Bernoulli blockmodel:

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{\theta_{11} L_{11}(\mathbf{x}) + \theta_{12} L_{12}(\mathbf{x}) + \theta_{21} L_{21}(\mathbf{x}) + \theta_{22} L_{22}(\mathbf{x})\}$$

c. Homogeneous Dyad-independent:

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{\theta L(\mathbf{x}) + \rho M(\mathbf{x})\}$$

d. Dyad-independent block

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{\sum_{r,s=1,2} \theta_{rs} L_{rs}(\mathbf{x}) + \rho M(\mathbf{x})\}$$

e. Dyad-independent block with block reciprocity parameters:

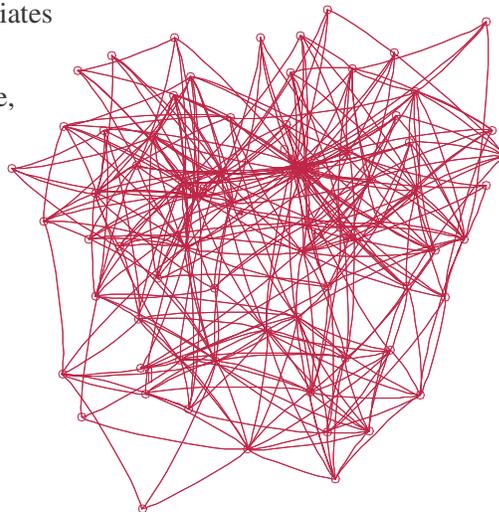
$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{\sum_{r,s=1,2} (\theta_{rs} L_{rs}(\mathbf{x}) + \rho_{rs} M_{rs}(\mathbf{x}))\}$$

Collaboration network in a US law firm (Lazega)

36 partners, 35 associates

3 offices

2 practices (corporate,
litigation)



Parameter estimates for Dyad-independent block model:

Lazega lawyer data

Block 1 = partners; block 2 = associates

d. Dyad-independent block

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{\sum_{r,s=1,2} \theta_{rs} L_{rs}(\mathbf{x}) + \rho M(\mathbf{x})\}$$

<i>Parameter</i>	<i>Estimate</i>
θ_{11}	-1.34
θ_{12}	-3.51
θ_{21}	-1.53
θ_{22}	-1.98
ρ	1.52

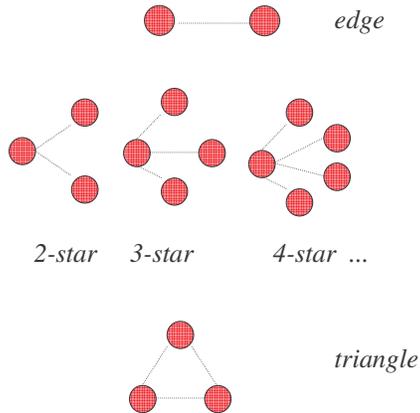
Dependence assumptions:

3. Markov random graphs

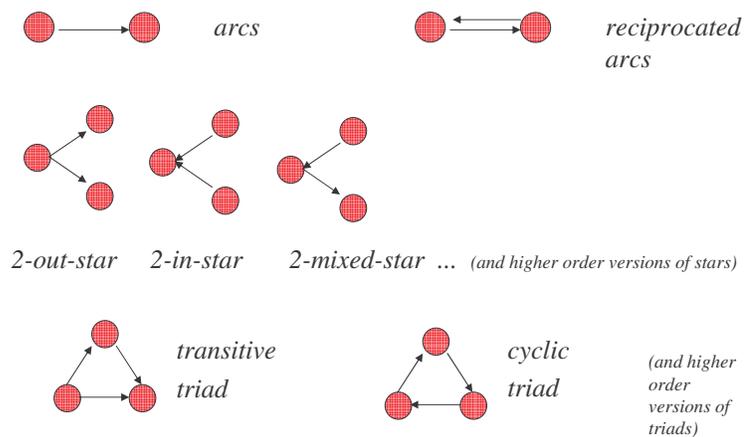
Suppose that edges are conditionally independent if and only if they share a node. (Frank & Strauss, 1986)

Frank and Strauss showed that configurations in this model comprised edges, stars and triangles.

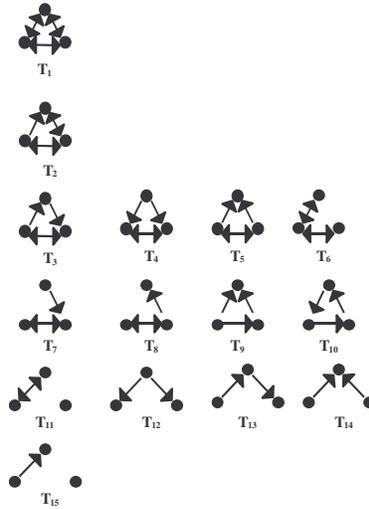
Network configurations for non-directed Markov random graph models



Some network configurations for directed Markov random graph models



**All triadic network configurations for directed
Markov random graph models**



**A Markov random graph model:
Nondirected networks**

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{\theta L + \sigma_2 S_2 + \sigma_3 S_3 + \tau T\}$$

Edge parameter (θ)

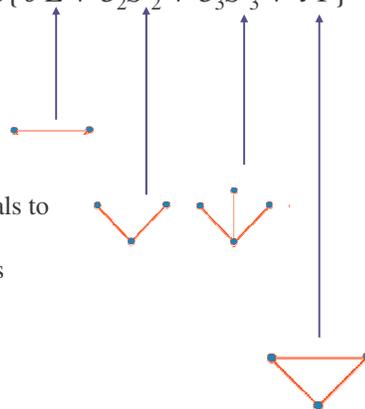
- L ... number of edges

Star parameters (σ_k)

- Propensities for individuals to have connections with multiple network partners

Triangle parameter (τ)

- represents clustering



Examples of graphs from Markov graph distributions

Bernoulli

$$\Pr(X = x) = (1/\kappa) \exp\{-2 L\}$$

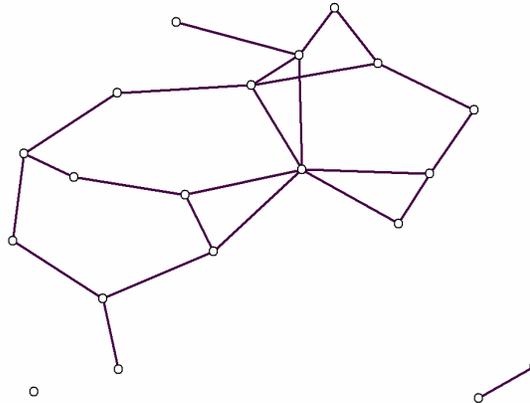
Graph statistics:

Edges = 24

2-stars = 51

3-stars = 34

Triangles = 3



Examples of graphs from Markov graph distributions

Markov: edge/star model

$$\Pr(X = x) = (1/\kappa) \exp\{-2 L + 0.5 S_2 - 0.3 S_3\}$$

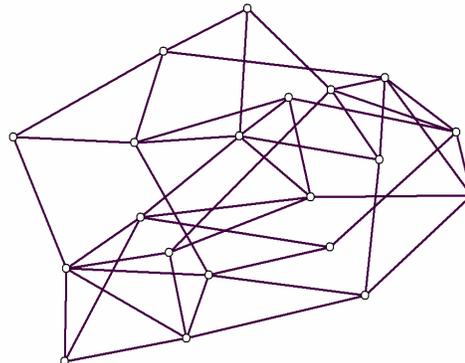
Graph statistics:

Edges = 44

2-stars = 158

3-stars = 148

Triangles = 12



Examples of graphs from Markov graph distributions

Markov with triangles

$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{-2L + 0.5S_2 - 0.3S_3 + T\}$$

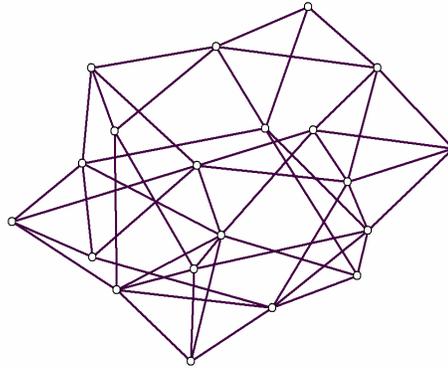
Graph statistics:

Edges = 49

2-stars = 201

3-stars = 224

Triangles = 21



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Simulating Markov random graph distributions

Using the *Metropolis algorithm*:

random starting graph on fixed number of nodes
many iterations (e.g. 1,000,000)
sample every 1000th graph

The algorithm sets up a Markov Chain on the space of all possible nondirected networks that has $\Pr(\mathbf{X}=\mathbf{x})$ as its stationary distribution

at each step, we consider changing the value of the (i,j) tie (from 1 to 0 or 0 to 1) for a randomly selected pair (and relation): $\mathbf{x} \rightarrow \mathbf{x}'_{ij}$
the change is made with probability

$$\min[1, \exp(\{\sum_Q \lambda_Q (z_Q(\mathbf{x}'_{ij}) - z_Q(\mathbf{x}))\})]$$

[burnin](#)

Estimation of Markov random graph models

1. **Pseudo-likelihood estimation**: An approximate technique based on logistic regression. Maybe indicates major effects but needs to be used with caution.
2. **Markov Chain Monte Carlo Maximum Likelihood Estimation (MCMCMLE)**: simulation of a distribution of random graphs from a starting set of parameter values, and subsequent refinement of the parameter values by comparing the distribution of graphs against the observed graph, with this process repeated until the parameter estimates stabilize.

[estimation](#)

Estimation software (MCMCMLE)

Statnet:

csde.washington.edu/statnet

SIENA:

stat.gamma.rug.nl/siena.html

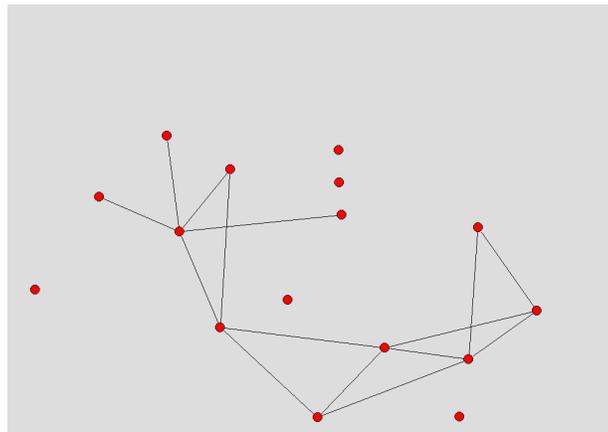
pnet:

www.sna.unimelb.edu.au/pnet/pnet.html

Example:

Florentine families business network

(Padgett & Ansell, 1993)



Markov model estimates: Florentine families business network

Model containing edges, 2-stars, 3-stars, triangles

Maximum Likelihood estimates

Edge = - 4.27 (1.13)*

2-star = 1.09 (0.65)

3-star = -0.67 (0.41)

Triangle= 1.32 (0.65)*

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Edge = - 4.27 (1.13)*

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Triangle= 1.32 (0.65)*

Goodness of fit: Florentine Business network

Comparing observed network to a sample from simulation of
Markov model: *t*-statistics

Model statistics

Edges: $t = -0.01$
 2-stars: $t = -0.01$
 3-stars: $t = 0.01$
 triangles: $t = -0.03$

Degree distribution

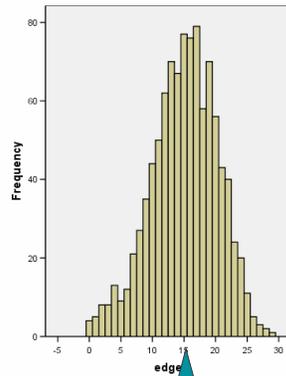
std dev degree dist: $t = 0.51$
 skew deg dist: $t = 0.17$

Clustering

global clustering: $t = 0.12$
 local clustering: $t = 0.68$

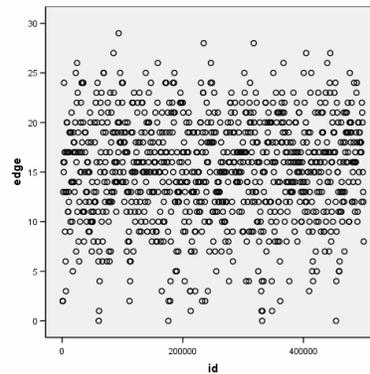
Geodesic distribution:

None of the quartiles of the geodesic distribution for the observed graph are extreme in the distribution

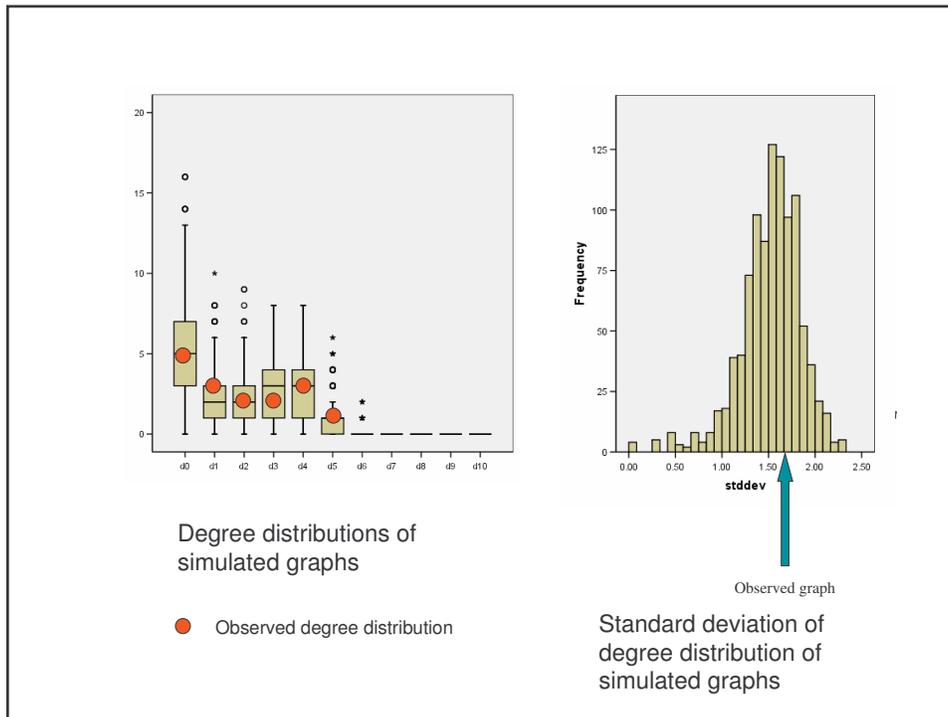


Observed graph

No of edges in simulated graphs



No of edges in simulated graphs against simulation number (*id*)



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Edge/Triangle model:

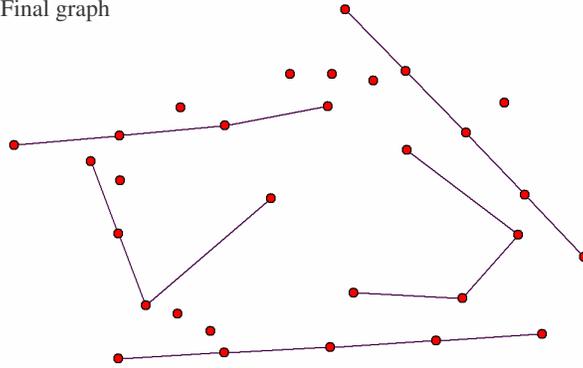
$\theta = -3$; $\tau = 0.25$

Start from **empty** graph

Burn-in 50,000

200,000 simulations

Final graph



Edge/Triangle model:

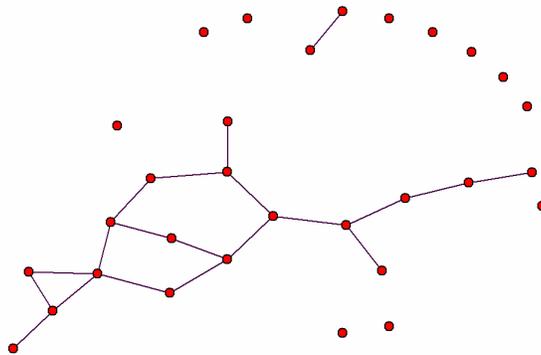
$\theta = -3$; $\tau = 0.75$

Start from **empty** graph

Burn-in 50,000

200,000 simulations

Final graph



Edge/Triangle model:

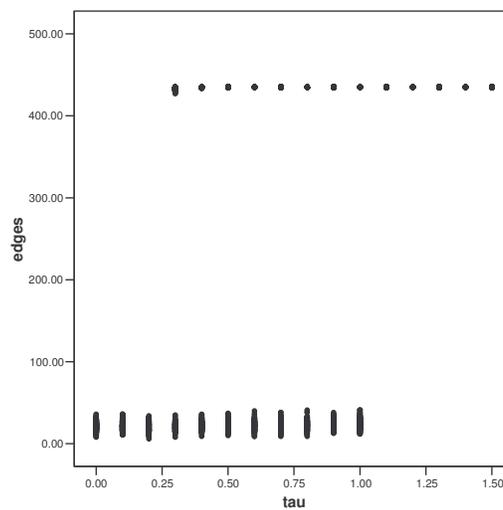
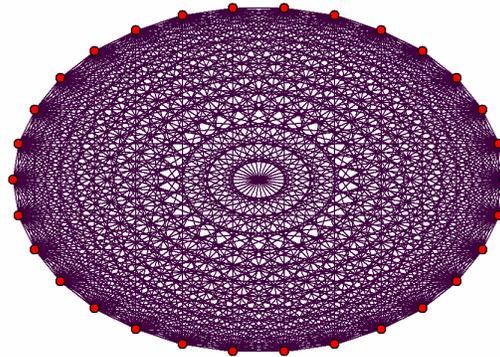
$\theta = -3$; $\tau = 0.75$

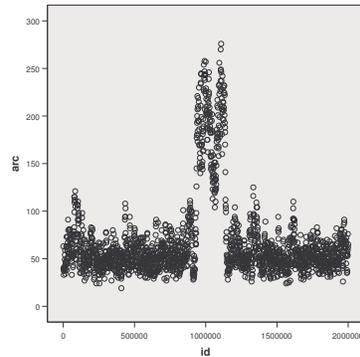
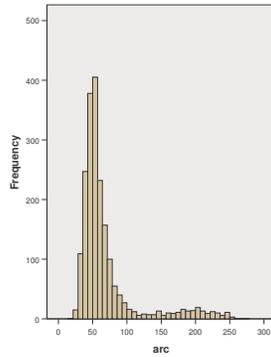
Start from **complete** graph

Burn-in 50,000

200,000 simulations

Final graph





What we DON'T want to see:
the model exhibits two regions!
This would indicate a bad model.

Model degeneracy

For certain parameter values, a model may imply that only one or two graphs with non-zero probability.

Often such graphs are the empty or full graph
(or a graph of complete disconnected components).

These *near degenerate* models cannot be estimated.

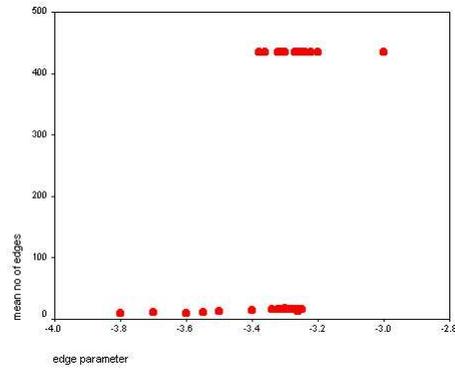
Markov models are often degenerate when clustering is high

Model degeneracy

Model with only an edge and triangle parameter $\Pr(\mathbf{X} = \mathbf{x}) = (1/c) \exp\{\theta L + \tau T\}$

Set $\tau = 1.0$

30 node graph



Model degeneracy

Model with edge, triangle parameter and star parameters

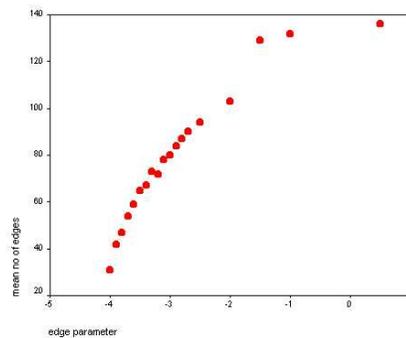
$$\Pr(\mathbf{X} = \mathbf{x}) = (1/c) \exp\{\theta L + \sigma_2 S_2 + \sigma_3 S_3 + \tau T\}$$

Set $\tau = 1.0$

$\sigma_2 = 0.5$

$\sigma_3 = -0.2$

30 node graph



Model degeneracy

For Markov random graphs, the bottom line is:

Edge/triangle models are NOT going to fit observed data:

They imply graphs that are either close to empty or close to complete (*degenerate*)

Models with edges, triangles, 2-star and 3-star parameters will do better:

They may fit data if there is a negative 3-star parameter estimate

But they are also likely to fail when clustering is high.

Model degeneracy

Why do Markov models fail when clustering is high in observed networks?

An attempt at an intuitive explanation:

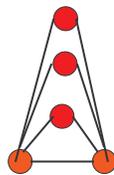
Markov models suppose that triangles are spread rather evenly throughout the graph (homogeneous models).

They have trouble when triangles tend to clump into *cohesive subsets of nodes*. (The dense part of the graph suggests a very strong triangle effect – but applying this effect to the not-so-dense part tends to “fill the graph up”.)

One solution

Go beyond Markov dependence assumptions:
Higher order dependence assumptions producing,
most importantly, *alternating k-triangle* statistics
(plus some other higher order statistics.)

3-triangle (T_3)



Snijders, Pattison, Robins & Handcock, 2006;

Overview

A. Introduction

- Why statistical models of social networks?
- The conceptual basis for exponential random graph models

B. Bernoulli, Dyadic independence and Markov models

- Simulation and estimation

C. Goodness of fit

D. Degeneracy

E. New specifications

- Degree sequences
- Higher order triangulation

F. Social selection models: incorporating attributes

Papers on new specifications

Advance copies: <http://www.sna.unimelb.edu.au/>

Original paper:

Snijders, Pattison, Robins, & Handcock (2006). New specifications for exponential random graph models. *Sociological Methodology*. In press.

Forthcoming in *Social Networks*:

Goodreau (2006). Advances in Exponential Random Graph (p^*) Models Applied to a Large Social Network.

Hunter (2006). Curved exponential family models for social networks.

Robins, Pattison, Kalish, & Lusher (2006). An introduction to exponential random graph (p^*) models for social networks.

Robins, Snijders, Wang, Handcock, & Pattison (2006). Recent developments in exponential random graph (p^*) models for social networks.

Also see Hunter & Handcock (2006). Inference in curved exponential family models for networks. *Journal of Computational and Graphical Statistics*. In press.

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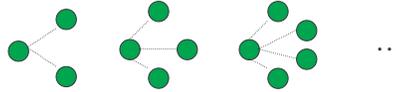
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Related model parameters

Star configurations



Parameters

σ_2 σ_3 σ_4 ...

Markov models with star effects only (ignore triangles)

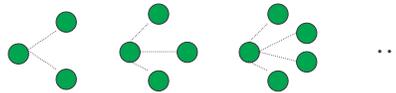
$$\Pr(\mathbf{X} = \mathbf{x}) = (1/\kappa) \exp\{ \theta L + \sigma_2 S_2 + \sigma_3 S_3 + \sigma_4 S_4 + \dots \}$$

We can combine all the star effects into the one parameter by setting constraints among the star parameters:

$$\sigma_2, \sigma_3, \sigma_4 \dots$$

Related model parameters

Star configurations



Parameters

σ_2 σ_3 σ_4 ...

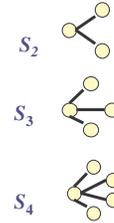
If we assume that $\sigma_k = -\sigma_{k-1}/\lambda$, for $k > 1$ and $\lambda \geq 1$ a (fixed) constant
alternating k-star hypothesis

Then we obtain a single *star* parameter (σ_2) with statistic:

$$S^{[\lambda]}(\mathbf{x}) = \sum_k (-1)^k S_k(\mathbf{x}) / \lambda^{k-2} \qquad \textit{alternating k-star statistic}$$

**1. Parameters for degree sequences:
a. alternating k -star parameters**

$$z(\mathbf{x}) = S_2 - \frac{S_3}{\lambda} + \frac{S_4}{\lambda^2} - \dots + (-1)^{n-2} \frac{S_{n-1}}{\lambda^{n-3}}$$

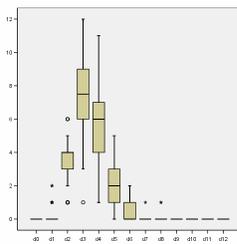


Interpretation:

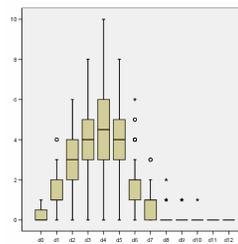
Positive parameter indicates centralization through a small number of high degree nodes
core-periphery based on popularity
More dispersed degree distribution

Negative parameter: “truncated” (less dispersed) degree distribution; nodes tend not to have particularly high degrees.

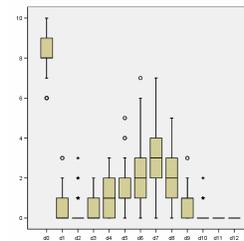
**Simulated degree distributions:
20 nodes: fixed density = 0.2**



Alt.kstar=-3



Alt.kstar=0



Alt.kstar=+3

Higher order models: **increasingly dispersed degree distribution**

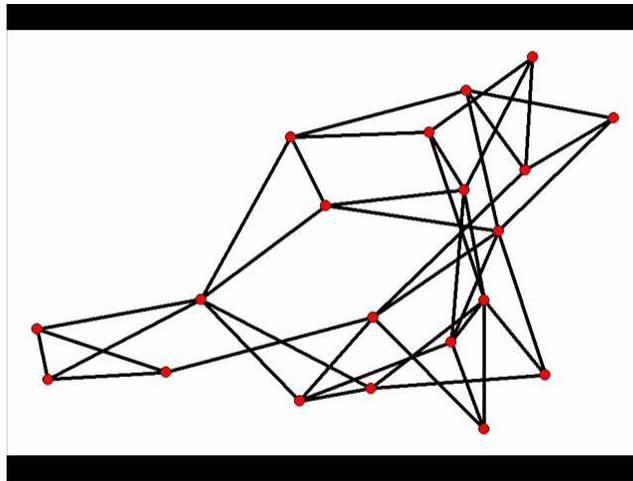
20 nodes: Fix the density at 0.20

all parameters = 0

EXCEPT

Vary *alternating k-star* parameter from -3.0 to $+3.0$
in steps of $+0.5$

The movie shows one representative graph from each simulated distribution



1. Parameters for degree sequences: b. Geometrically weighted degree distributions

An equivalent characterisation:

Consider statistics $d_k(\mathbf{y})$, where $d_k(\mathbf{y})$ is the number of nodes in \mathbf{y} of degree k (with corresponding parameters θ_k)

Assuming that $\theta_k = e^{-\alpha k \gamma}$ for $k = 1, 2, \dots, n-1$

yields the statistic:

$$D^{[\alpha]}(\mathbf{y}) = \sum_k e^{-\alpha k} d_k(\mathbf{y}) \quad \text{weighted degree distribution}$$

Relationship with alternating k-star statistic:

$$S^{[\lambda]}(\mathbf{y}) = \lambda^2 D^{[\alpha]}(\mathbf{y}) + 2\lambda L(\mathbf{y}) - n \lambda^2 \quad \text{See Hunter (2006) } \lambda = e^{\alpha/(e^{\alpha}-1)}$$

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- Higher order triangulation

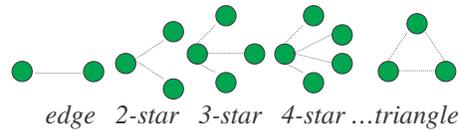
F. Social selection models: incorporating attributes

Neighborhoods depend on proximity assumptions

Assumptions: two ties are neighbors: Configurations for neighborhoods

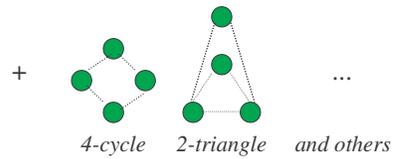
if they share an actor

Markov



if they complete a 4-cycle

realization-dependent



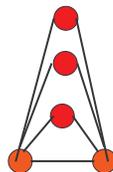
2. Parameters for higher order triangulation

Cohesive subsets of nodes

Require realization dependent neighbourhoods:

Alternating k -triangles:

3-triangle (T_3)



2. Parameters for higher order triangulation

Alternating k -triangles:

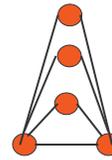
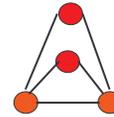
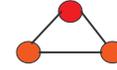
$$u(\mathbf{y}) = T_1 - \frac{T_2}{\lambda} + \frac{T_3}{\lambda^2} - \dots + (-1)^{n-2} \frac{T_{n-2}}{\lambda^{n-3}}$$

Interpretation:

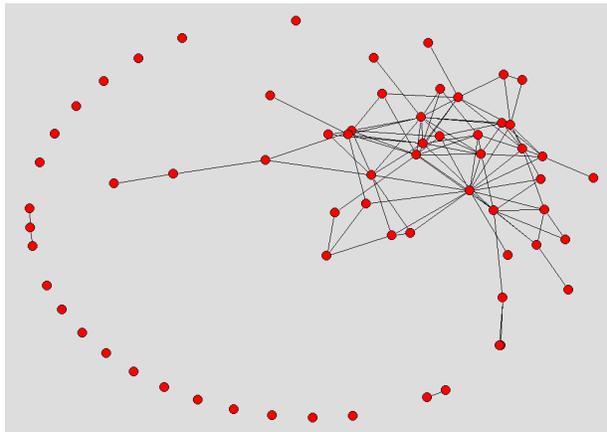
a. Positive parameter suggests triangles tend to "clump" together in denser regions of the network (cohesive subsets).

b. Models the **edgewise shared partner distribution**:

For each pair of tied nodes, how many partners do they share? (Hunter, 2006)



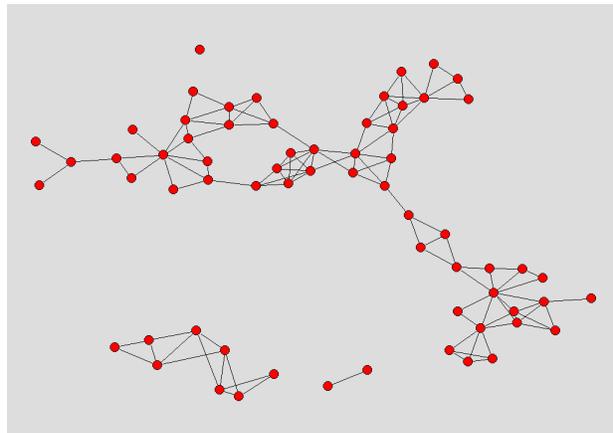
etc...



Interpretation:
With **only** positive k -triangle effect, there is a core-periphery type structure based on triangulation

Higher order model
Positive k -triangle parameter:

Parameters:
Edge = -4.5
Alt. k -triangle=1.3
65 nodes



Interpretation:
 With positive k -triangle effect, and negative k -star, various regions of greater density distributed across the network.

Higher order model

Positive k -triangle parameter & negative k -star parameter:

Parameters:
 Edge = -0.5, alt. k -star = -1.5, alt. k -triangle = 2.0
 65 nodes

Higher order models: **increasing triangulation**

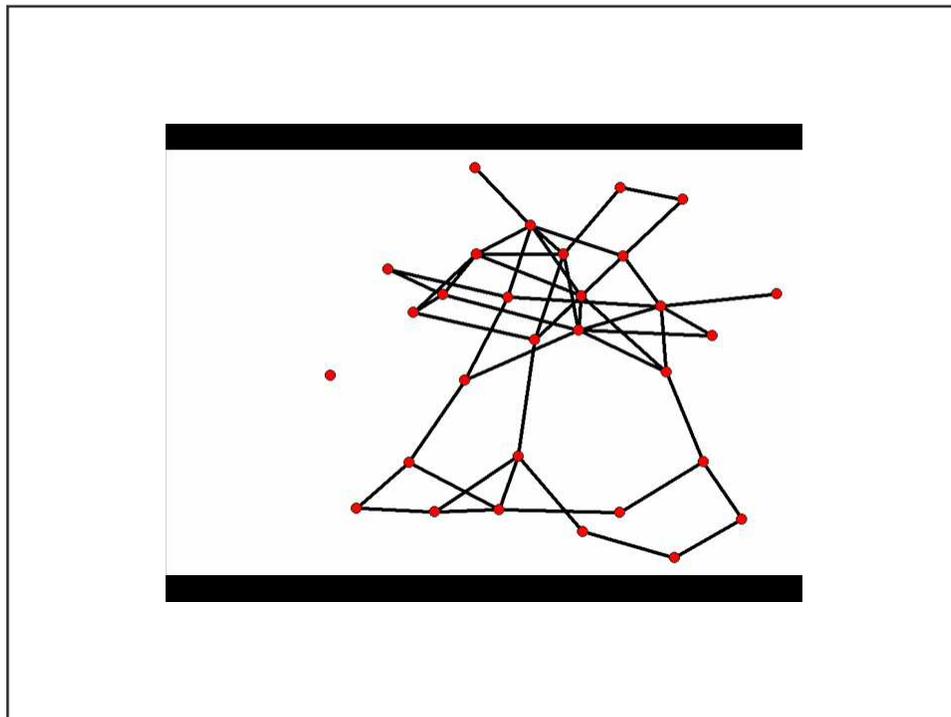
30 nodes: Fix the density at 0.10

all parameters = 0

EXCEPT

Vary *alternating k -triangle* parameter from 0.0 to +3.0
 in steps of +0.25

The movie shows one representative graph from each simulated distribution



Higher order models:

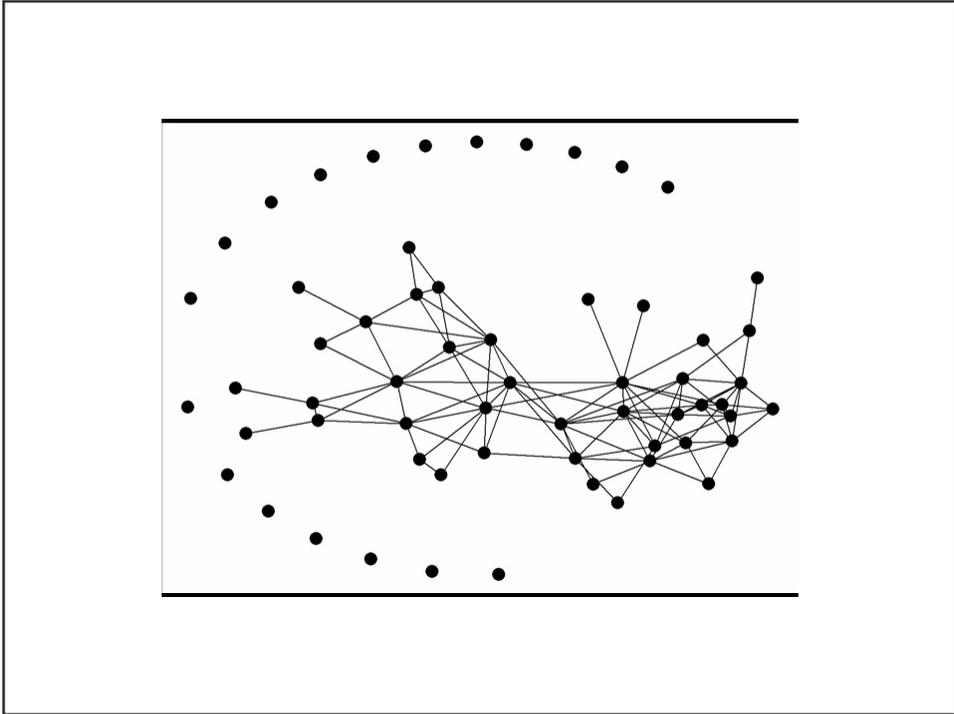
From centralization to segmentation

60 nodes: Fix the density at 0.05 (i.e. 88 or 89 edges)

alternating k-triangle parameter = +2.0

Vary *alternating k-star* parameter from 0.0 to - 1.0
in steps of - 0.1

The movie shows one representative graph from each
simulated distribution

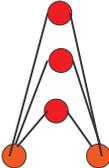


3. Parameters for multiple connectivity

Alternating independent 2-paths:

3-indept.2-paths
(U_3)

$$z(\mathbf{x}) = U_1 - \frac{2U_2}{\lambda} + \sum_{k=3}^{n-2} \left(\frac{-1}{\lambda}\right)^{k-1} U_k$$



Models the **dyadwise shared partner distribution**:

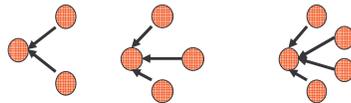
For each pair of nodes, how many partners do they share? (Hunter, 2006)

**Directed networks:
Alternating k -star parameters**

Snijders, Pattison, Robins & Handcock (2006)

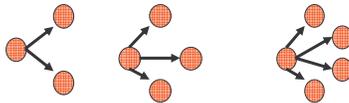
Alternating k -in-stars:

Statistic: Weighted summation of k -in-stars analogous to the non-directed model



Alternating k -out-stars:

Statistic: Weighted summation of k -out-stars analogous to the non-directed model

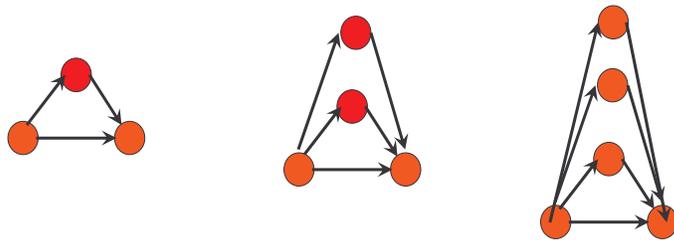


Directed networks:

Snijders, Pattison, Robins & Handcock (2006)

Alternating directed k -triangles:

Statistic: weighted sum of directed k -triangles with 2-paths as sides of the k -triangle

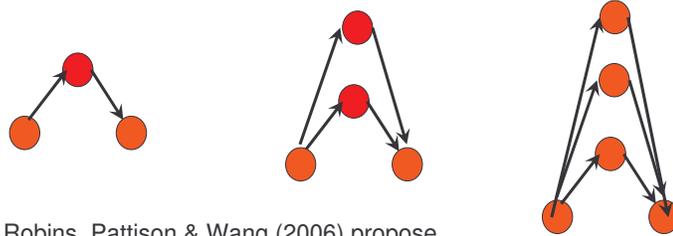


Directed network parameters

Snijders, Pattison, Robins & Hancock (2006)

Alternating directed k -2paths:

Statistic: weighted sum of k -2paths



NB: Robins, Pattison & Wang (2006) propose additional variations for directed network models

Fitting Models: 20 network data sets from UCINET5

(Borgatti, Everett & Freeman, 1999)

Non directed networks:

Kapferer mine: kapfmm, kapfmu (16 nodes)
Kapferer tailor shop: kapfts1, kapfts2 (39 nodes)
Padgett Florentine families: padgb, padgm (16 nodes)
Read Highland tribes: gamapos (16 nodes)
Zachary karate club: Zache (34 nodes)
Bank wiring room: rdpos, rdgam (14 nodes)
Taro exchange: Taro (22 nodes)
Thurman office: Thurm (15 nodes)

Directed networks:

Kapferer tailor shop: kapfti1, kapfti2 (39 nodes)
Wolf primates: wolfk (20 nodes)
Krackhardt hi-tech managers: friend, advice (21 nodes)
Bank wiring room: rdhlp (14 nodes)
Knoke bureaucracies: knokm, knoki (10 nodes)

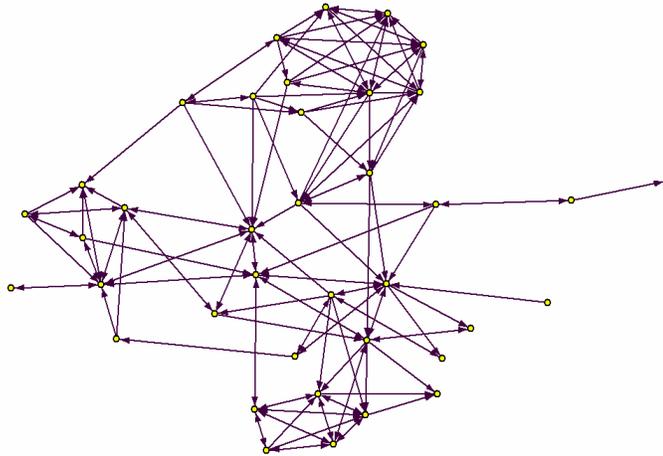
Non directed networks

Data set	Markov	Higher order
Kapfmm	OK	OK
Kapfmu	OK	OK
Kapfts1	Does not converge	OK
Kapfts2	Does not converge	OK
Padgm	OK	OK
Padgb	OK	OK
Gamapos	Does not converge	OK
Zache	Does not converge	OK
Rdpos	OK	OK
Rdgam	OK	OK
Taro	Does not converge	OK
Thurm	OK	OK
TOTAL	7/12	12/12

Directed networks

Data set	Markov	Higher order
Kapfti1	Does not converge	OK
Kapfti2	Does not converge	OK
Wolfk	Does not converge	OK
Krackhardt friend	OK	OK
Krackhardt advice	OK	OK
Rdhlp	OK	OK
Knokm	OK	OK
Knoki	OK	OK
TOTAL (directed)	5/8	
TOTAL (nondir.)	7/12	
TOTAL	12/20	20/20

Example: After hours socialising network



After hours network

<u>Parameter</u>	<u>Estimate</u>	<u>Standard error</u>	<u>Convergence statistic</u>
Arc	- 1.27*	0.63	0.06
Reciprocity	2.42*	0.34	0.06
<i>k</i> instar	- 0.86*	0.32	0.06
<i>k</i> outstar	- 0.96*	0.33	0.06
Alt. <i>k</i> -triangles	1.09*	0.14	0.06

Goodness of fit: After hours network

Model statistics

Arcs: $t = 0.03$
Reciprocity: $t = 0.03$
 k instar (2) : $t = 0.03$
 k outstar (2) : $t = 0.02$
Alt. k -triangles (2): $t=0.03$

Other Markov statistics

2-in-stars: $t = - 0.14$
3-in-stars : $t = - 0.37$
2-out-stars: $t = - 0.29$
3-out-stars : $t = -0.53$
2-paths: $t = - 0.34$
Transitive triads: $t = - 0.07$
Cyclic triads: $t = - 0.37$

Other higher order statistics

Alt. 2-paths (2): $t = - 0.37$

Goodness of fit: After hours network

Degree distributions

standard deviations
indegrees: $t = - 0.10$
outdegrees : $t = - 0.56$
skew
indegrees: $t = - 1.05$
outdegrees : $t = - 1.96$

Clustering coefficients

Proportion of 2-stars in transitive triads:
2-in-stars: $t = 0.40$
2-out-stars: $t = 1.18$
2-paths: $t = 1.57$
Proportion of 2-paths
in cyclic triads: $t = -0.13$

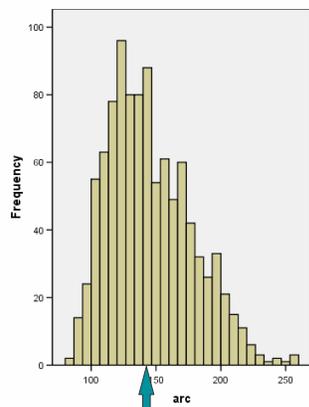
Geodesic distribution:

None of the quartiles of the geodesic distribution for the observed graph are extreme in the distribution

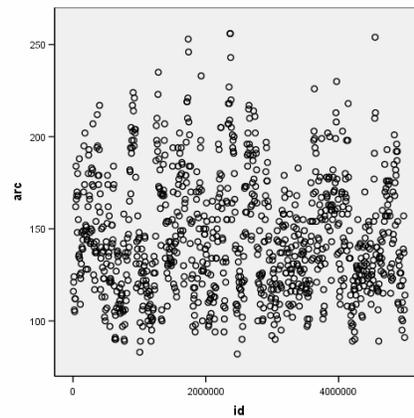
Goodness of fit: After hours network

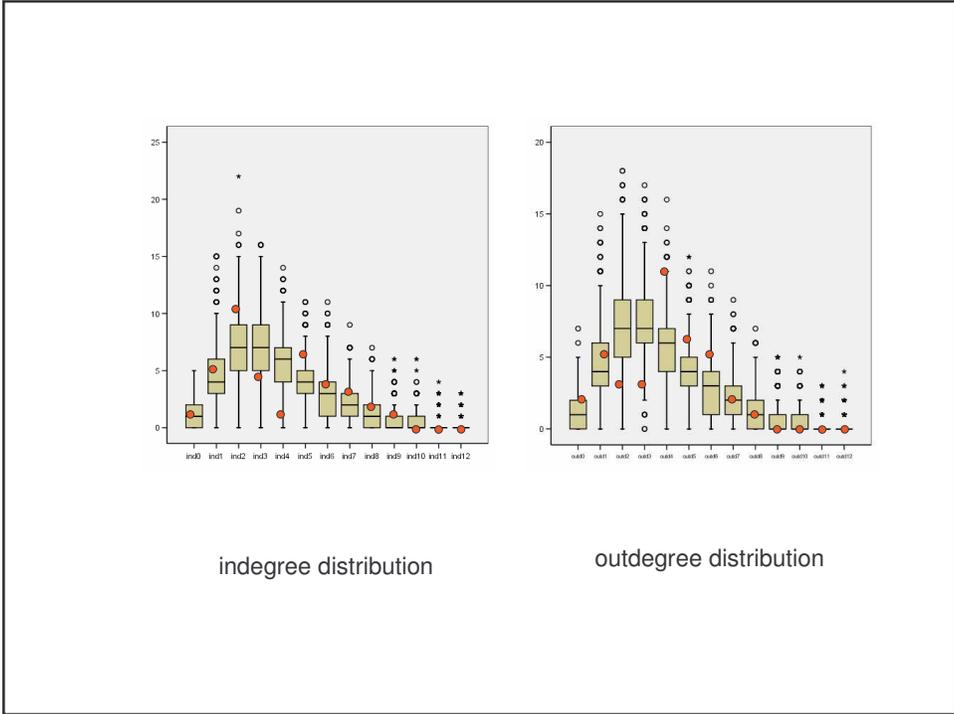
Triad census

300: $t = -0.5036$
210: $t = -0.1364$
120C: $t = -0.3707$
120D: $t = 1.4142$
120U: $t = 1.0437$
201: $t = -0.4520$
111D: $t = -0.1218$
111U: $t = -0.5027$
030T: $t = 2.2584$
030C: $t = -0.3287$
102: $t = 0.4598$
021D: $t = -0.4971$
021C: $t = -0.9884$
021U: $t = 0.1030$
012: $t = 0.3200$
003: $t = -0.1781$



Observed graph





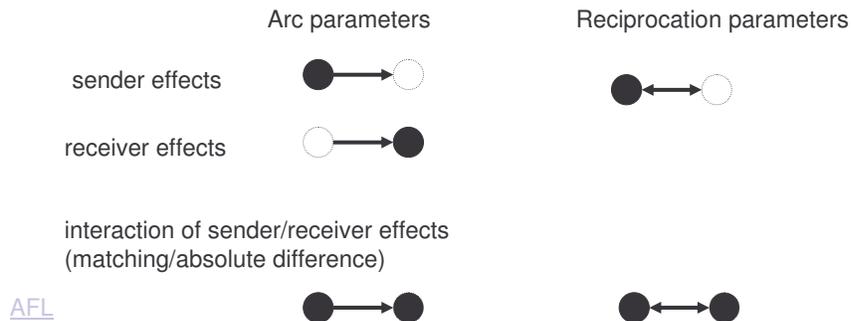
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Models with nodal attributes: Social selection models

There are various ways to introduce actor attributes (binary or continuous) Robins, Elliott & Pattison (2001)

e.g. dyad level effects



Concluding remarks

- The new specifications are a dramatic improvement over the previous models. For many data sets we now have models that are quite convincing in reproducing major features of the network.
 - Larger networks are more difficult to fit; Directed networks are more difficult to fit.
- This is NOT to say all problems are solved
 - Degeneracy may still be a problem for these models applied to certain data sets (indicates the model specification is not right for that data.)
- So ongoing work required on model specification, BUT
 - Realization dependence models give us a substantial way forward.
 - Combining related parameters into the one function through weighted constraints is parsimonious and helps with model convergence
- MCMCMLE methods of parameter estimation, and model simulation techniques are a crucial part of these recent developments

Work in progress

- Directed networks
- Bipartite graphs (affiliation networks)
 - Also tripartite graphs
- Multiple networks
- Longitudinal models

To come:

- Missing data
- Network sampling