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# The prognostic value of phosphorylated Akt in breast cancer: a systematic review

SUBJECT AREAS:  
PROGNOSTIC MARKERS  
BREAST CANCER

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The prognostic value of phosphorylated Akt (pAkt) overexpression in breast cancer has been investigated by many studies with inconsistent results. This systematic review was conducted to evaluate the association of pAkt overexpression with breast cancer prognosis in terms of overall survival and disease-free survival. Three electronic databases (PubMed, EMBASE and Chinese Biomedical Literature Database) were comprehensively searched. Hazard ratios (HRs) with 95% confidence intervals (CIs) from different studies were combined using the random-effects model. In total, 33 studies with 9,836 patients were included for final analysis. The summary HR for overall survival and disease-free survival was 1.52 (95% CI: 1.29–1.78) and 1.28 (95% CI: 1.13–1.45), respectively, indicating higher risk of death and disease recurrence associated with pAkt overexpression. The results were robust in sensitivity analyses by omitting one study each time and by using the fixed-effects model instead. Subgroup and meta-regression analyses did not show that the prognostic effect of pAkt overexpression would change materially with such factors as population, status of hormone receptors, hormonal or trastuzumab treatment given, analyzing method (univariate versus multivariate) and methodological quality of the original studies. In conclusion, the available evidence suggests that pAkt overexpression is an adverse prognostic factor for breast cancer.

**B**reast cancer has long been the most frequent cancer among women worldwide, with an estimated 1.67 million new cases diagnosed each year (25% of all cancers)<sup>1</sup>. Despite the significant progress in early detection and treatment over the past decades, breast cancer remains the leading cause of cancer deaths in women in many countries especially the less developed ones<sup>1</sup>. To achieve better management of breast cancer, the identification of clinical, pathological and biological factors that have prognostic value is very important, as those factors could be used to inform risk stratification, treatment selection and development of new therapeutic strategies<sup>2</sup>. Examples of such factors include tumor size, lymph node status, estrogen receptor (ER) status and human epidermal growth factor receptor 2 (HER2) status, which have been well integrated into clinical practice and contributed much to the improvement of breast cancer prognosis. Along with the emphasis on personalized medicine in recent years, increasing attention has been drawn to other biomarkers that may help explain residual risk not accounted for by the aforementioned traditional factors<sup>2</sup>.

Akt, also known as protein kinase B, is a serine/threonine protein kinase that, once activated by phosphorylation at serine 473 and threonine 308, plays an important role in multiple cellular processes<sup>3</sup>. In particular, phosphorylated Akt (pAkt) may induce signals interfering with the apoptotic functions of the cell, and promote cell survival, proliferation and motility possibly through activation of mammalian target of rapamycin among other mechanisms<sup>3–6</sup>. Overexpressed pAkt is frequently observed in human lung, gastric, hepatocellular, pancreatic, renal, prostate and endometrial cancer as well as multiple myeloma<sup>7–11</sup>. Studies have documented the prognostic role of pAkt overexpression in some cancers. For example, a recent meta-analysis showed that pAkt overexpression was significantly associated with worse overall survival in non-small cell lung cancer patients (hazard ratio [HR]: 1.38, 95% confidence interval [CI]: 1.11–1.70)<sup>12</sup>.

In breast cancer, the prognostic impact of this biomarker has also been evaluated by many studies, but their results were inconsistent. For example, the study of Xia et al with 130 patients found that pAkt overexpression was significantly associated with worse overall survival (HR: 2.16, 95% CI: 1.22–3.81)<sup>13</sup>. However, in the study of Fabi et al with 73 patients, no significant association between pAkt status and overall survival was found ( $P = 0.97$ )<sup>14</sup>.



The discrepancy between individual studies could have been due to multiple reasons such as different populations, sample sizes, methodological problems, and other potential confounding factors. Against this background, we conducted a comprehensive systematic review with an aim to clarify the prognostic value of pAkt overexpression in breast cancer. The potential impact of various factors on pAkt's prognostic effect was also investigated.

## Results

**Study selection and characteristics.** The flow of study selection is shown in Figure 1. Initially, 2,976 records, including 1,063 duplicates, were identified by our literature search. Among the 1,913 unique records, 173 studies were subject to full text examination and 33 studies were considered eligible and finally included for the present systematic review<sup>6,13–44</sup>. Two studies<sup>33,34</sup> were based on a same cohort with focus on different outcomes. The characteristics of the 33 studies are summarized in Table 1. Their sample sizes ranged from 44 to 1,355, with a median of 142. In total, 9,836 patients were included for analysis. All studies assessed pAkt status by immunohistochemistry, and most of them used mouse anti-pAkt (Ser473) antibodies. pAkt overexpression was found in 12.7% to 87.5% of the subjects, with a summary rate of 49.3% (95% CI: 42.4%–56.2%). Four and three studies clearly reported that all their

subjects received trastuzumab and hormone treatment, respectively, while the other studies made no clear statement on this issue. The study quality scores based on the 9-point Newcastle-Ottawa scale ranged from 5 to 9, with a median of 7 and a mean of 6.3.

**Meta-analyses.** HRs for overall survival were available from 20 of the 33 included studies (Table 1). Meta-analysis of the 20 studies with 6,349 patients showed that pAkt overexpression was significantly associated with worse overall survival in breast cancer. The summary HR was 1.52 (95% CI: 1.29–1.78), with substantial heterogeneity among the studies ( $I^2 = 58.4\%$ ,  $P = 0.001$ ) (Figure 2). Sensitivity analyses omitting one study each time showed that individually Wang 2009 and Wang C 2011 had the largest influence on the result. The summary HR became 1.46 (95% CI: 1.26–1.71; heterogeneity test  $I^2 = 54.6\%$ ,  $P = 0.002$ ) when Wang 2009 was omitted and 1.57 (95% CI: 1.32–1.87; heterogeneity test  $I^2 = 58.8\%$ ,  $P = 0.001$ ) when Wang C 2011 was omitted. When the fixed-effects model was used instead of the random-effects model, the summary HR became 1.31 (95% CI: 1.20–1.43).

HRs for disease-free survival were available from 24 of the 33 included studies (Table 1). Meta-analysis of the 24 studies with 8,683 patients showed that pAkt overexpression was significantly associated with worse disease-free survival in breast cancer. The summary HR was 1.28 (95% CI: 1.13–1.45), with substantial heterogeneity among the studies ( $I^2 = 74.2\%$ ,  $P < 0.001$ ) (Figure 3). Sensitivity analyses omitting one study each time showed that individually Aleskandarany 2011 and Yamamoto 2006 had the largest influence on the result. The summary HR became 1.33 (95% CI: 1.15–1.53; heterogeneity test  $I^2 = 68.1\%$ ,  $P < 0.001$ ) when Aleskandarany 2011 was omitted and 1.24 (95% CI: 1.10–1.40; heterogeneity test  $I^2 = 72.2\%$ ,  $P < 0.001$ ) when Yamamoto 2006 was omitted. When the fixed-effects model was used instead of the random-effects model, the summary HR became 1.06 (95% CI: 1.02–1.10).

To investigate the heterogeneity detected in the above meta-analyses, a series of subgroup and meta-regression analyses were conducted as planned. Although the summary HRs were not statistically significant in some subgroups (e.g. the one with proportion of ER-positive patients  $< 50\%$ ; the one in which trastuzumab was given to all patients), meta-regression analyses suggested that the between-subgroup differences did not reach statistical significance (Table 2). More rigorous stratification of studies according to the stage of cancer, ER status, PR status, and HER2 status did not show significant difference between the subgroups either (for details, see Supplementary Table S1). Thus, there is no evidence to show that any of these factors could explain the heterogeneity. In other words, the prognostic effect of pAkt overexpression did not change materially with such factors as population, sample size, status of hormone receptors, and methodological features of the original studies.

**Analysis of Publication Bias.** The funnel plots corresponding to Figure 2 and Figure 3 demonstrated some degree of asymmetry (Egger's regression tests:  $P < 0.001$  and  $P = 0.009$ , respectively) (Figure 4), which could be due to potential publication bias among other reasons<sup>45</sup>. After adjusting for the potential publication bias by trim-and-fill method, the summary HRs corresponding to Figure 2 and Figure 3 became 1.35 (95% CI: 1.15–1.58) and 1.22 (95% CI: 1.08–1.39), respectively. Although the adjusted estimates were slightly smaller than the unadjusted ones, they were still statistically significant and did not influence the original conclusion.

## Discussion

The present systematic review included 33 studies with 9,836 patients to evaluate the prognostic effect of pAkt overexpression in breast cancer, representing the most comprehensive summary of available

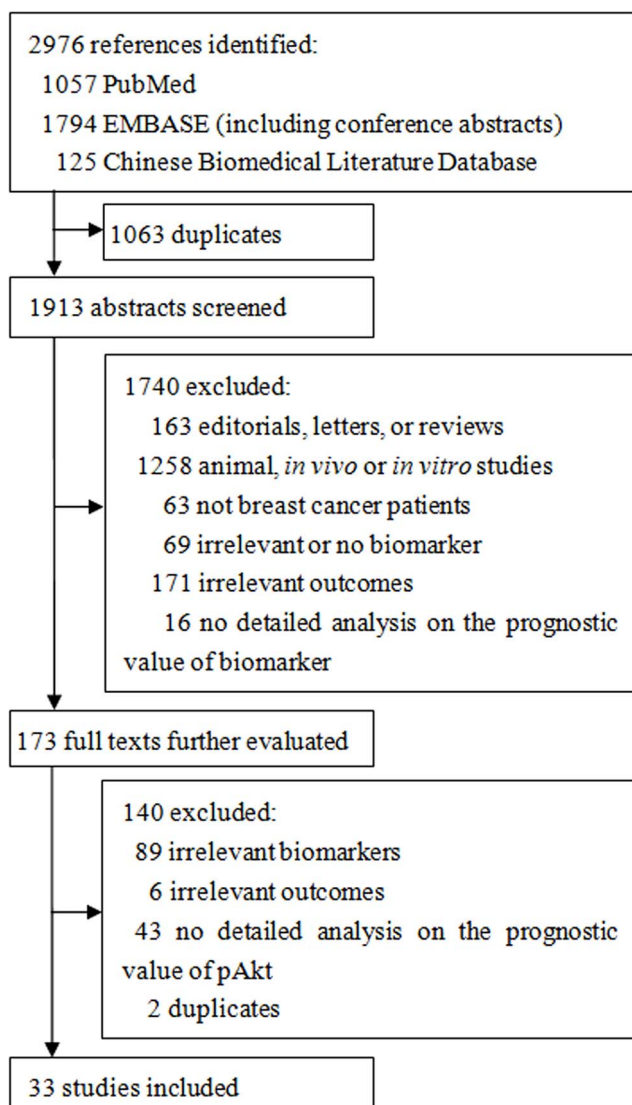


Figure 1 | Flow chart of study selection.



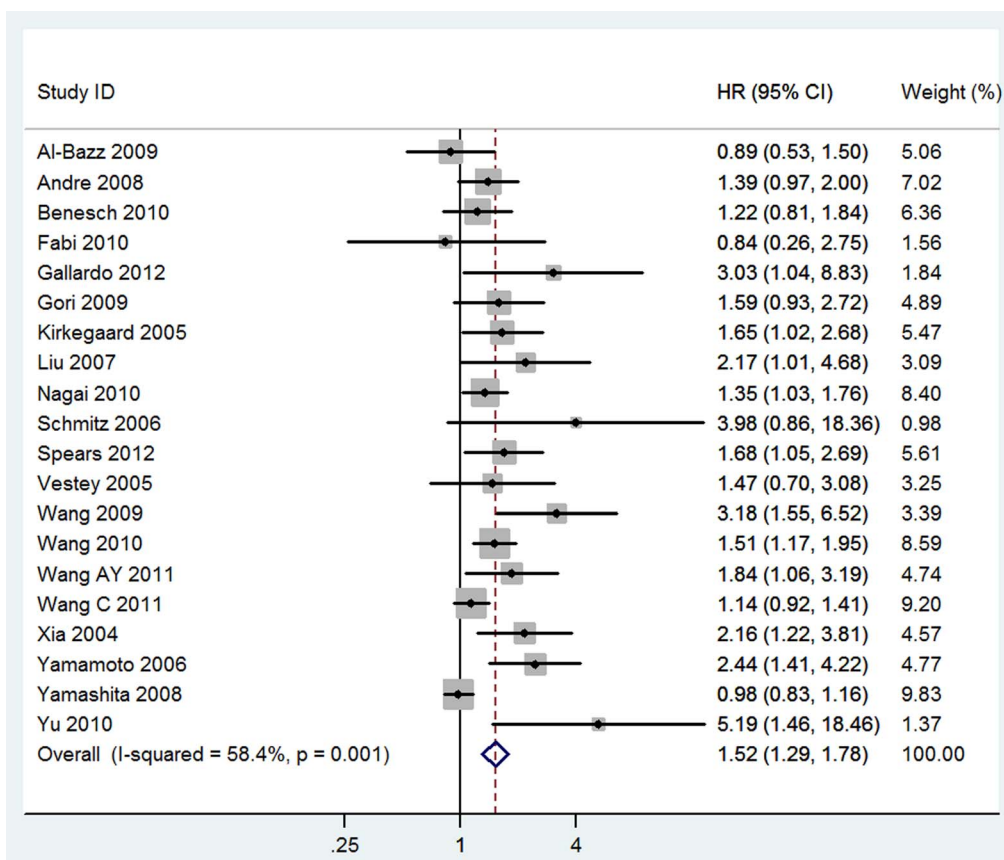
Table 1 | Characteristics of included studies

Study	Country	Period	N	Stage	ER + (%)	PR + (%)	HER2 + (%)	pAkt + (%)	pAkt detection method	Treatment	Mean FU (year)	Outcome	HR	NOS
AlBazz 2009	UK	1994~1997	106	--	59	40	--	21.7	IHC, Ser473	S	5.0	OS, DFS	Uni	5
Aleksandarany 2011	UK	1990~1998	1202	1~3	69	56	13	76.0	IHC, Ser473	S	20.0	DFS	Uni	7
An 2010	Korea	1992~2006	560	1~3	57	49	24	55.2	IHC, Ser473	S	4.9	DFS	Multi	7
Andre 2008	France	1989~1995	752	1~3	86	76	13	15.2	IHC	S	10.0	OS, DFS	Multi	8
Benesch 2010	Germany	1985~1995	160	--	38	52	--	58.1	IHC	S	≥13.0	OS	Uni	6
Capodanno 2009	Italy	1988~1998	72	--	63	--	--	87.5	IHC, Ser473	S/C/R/H	10.0	DFS	Uni	8
Cicenas 2005	Switzerland	--~1996	156	--	74	55	39	13.5	IHC, Ser473	--	4.8	DFS	Multi	8
Fabi 2010	Italy	2004~2007	73	4	39	36	100	71.2	IHC, Ser473	C/T	2.0	OS	Uni	5
Gallardo 2012	Spain	--	143	1~4	--	--	100	28.0	IHC	C/T	5.3	OS	Uni	5
Gori 2009	Italy	1999~2006	45	4	49	42	100	51.1	IHC	S/T	1.9	OS	Uni	5
Hartog 2011	Netherlands	1996~2005	429	1~3	76	63	7	12.7	IHC, Ser473	C	4.6	DFS	Uni	5
Janssen 2007	Norwegian	1978~1994	125	1~3	53	62	52	43.9	IHC, Ser473	S	11.0	DFS	Uni	7
Kirkegaard 2005	UK	1983~1999	392	--	100	--	45	50.5	IHC, Ser473	C	6.5	OS	Uni	6
Liu 2007	China	1996~2000	130	1~3	100	--	31	46.9	IHC, Ser473	H	5.1	OS, DFS	Multi	8
Nagai 2010	Brazil	--	1026	1~4	67	45	14	48.1	IHC, Ser473	S	10.0	OS, DFS	Uni	6
Perez-Tenorio 2002	Sweden	1984~1996	93	--	76	93	7	53.8	IHC, Ser473	S/H	5.3	DFS	Multi	8
Schmitz 2006	Germany	1989~1996	113	1~3	62	39	26	64.6	IHC, Ser473	S	7.0	OS	Multi	8
Spears 2012	UK	1981~1998	1355	1~3	80	83	13	50.5	IHC, Thr308	S/R	5.0	OS, DFS	Multi	8
Sun 2006	China	1994~1998	260	1~3	--	--	--	50.0	IHC	S	5.0	DFS	Uni	7
Tokunaga 2006	Japan	1991~2002	240	1~3	64	46	25	66.3	IHC, Ser473	S	12.5	DFS	Multi	5
Vestey 2005	UK	1996~2000	95	--	64	--	84	81.1	IHC, Ser473	S	4.3	OS, DFS	Multi	5
Wang 2009, Wang XL 2011	China	1997~2007	110	--	46	46	32	40.9	IHC, Ser473	S	10.0	OS, DFS	Multi*	7
Wang 2010	China	2001~2005	97	1~3	--	--	--	77.3	IHC, Thr308	S	6.5	OS	Uni	6
Wang C 2011	Canada	--	944	1~2	81	--	10	46.7	IHC, Ser473	S	10.4	OS, DFS	Multi	8
Wang AY 2011	China	2001~2005	81	1~3	78	--	--	27.2	IHC	S	>5.0	OS, DFS	Uni	6
Wu 2008	US	1999~2005	141	1~4	--	--	33	50.4	IHC, Ser473	S/C	4.0	DFS	Multi	9
Xia 2004	China	1988~1994	130	1~3	43	50	68	26.2	IHC, Thr308	S	4.0	OS	Uni	7
Yamamoto 2006	Japan	1987~2002	221	1~4	--	--	--	41.2	IHC, Ser473	S/C	5.7	OS, DFS	Multi*	7
Yamashita 2008	Japan	1982~1999	278	--	100	--	23	--	IHC, Ser473	S	8.0	OS, DFS	Uni	5
Yonemori 2009	Japan	1999~2006	44	2~3	11	7	100	79.5	IHC	S/C//T	>5.0	DFS	Uni	5
Yu 2010	China	2003~2007	98	1~3	60	56	48	37.8	IHC, Ser473	S	3.0	OS, DFS	Uni	5
Zhou 2004	China	1988~1991	165	--	50	--	32	73.9	IHC, Ser473	S	6.4	DFS	Uni	7

Abbreviations: N = number of patients included for this meta-analysis; ER = estrogen receptor; PR = progesterone receptor; HER2 = human epidermal growth factor receptor 2; FU = follow-up; HR = hazard ratio; NOS = Newcastle-Ottawa scale; IHC = immunohistochemistry; S = surgery; OS = overall survival; DFS = disease-free survival; Uni = univariate; Multi = multivariate; C = chemotherapy; R = radiotherapy; H = hormonal therapy; T = trastuzumab.

\*The analysis on overall survival (Wang 2009) was univariate, while the analysis on disease-free survival (Wang XL 2011) was multivariate.

\*\*The analysis on overall survival was univariate, while the analysis on disease-free survival was multivariate.



**Figure 2 | Meta-analysis of the association between pAkt overexpression and overall survival in breast cancer.** Results are presented as individual and pooled HRs with corresponding 95% CIs. HR > 1 means that overall survival of the patients with pAkt overexpression is worse than that of the pAkt-negative ones, while HR < 1 means the opposite.

evidence on this topic so far. pAkt overexpression was found to be associated with both worse overall survival (HR: 1.52, 95% CI: 1.29–1.78) and worse disease-free survival (HR: 1.28, 95% CI: 1.13–1.45) in breast cancer. Specifically, pAkt-overexpressed patients have a 50% higher risk of death and a 30% higher risk of disease recurrence compared with those without pAkt overexpression.

Substantial between-study heterogeneity was detected in our meta-analyses. However, subgroup and meta-regression analyses provided no evidence that any of the pre-specified factors such as population, status of hormone receptors, hormonal or trastuzumab treatment given, effect measure used in the original studies (HR vs. rate ratio), analyzing method (univariate vs multivariate) and study quality accounted for the heterogeneity. On one hand, this indicated that the prognostic effect of pAkt overexpression was robust in various scenarios, while on the other hand we can infer that there were other factors than the investigated ones existing as effect modifiers. Based on the data collected, we suggested that the varying scoring methods for pAkt status and definitions of pAkt overexpression have at least partly contributed to the between-study heterogeneity.

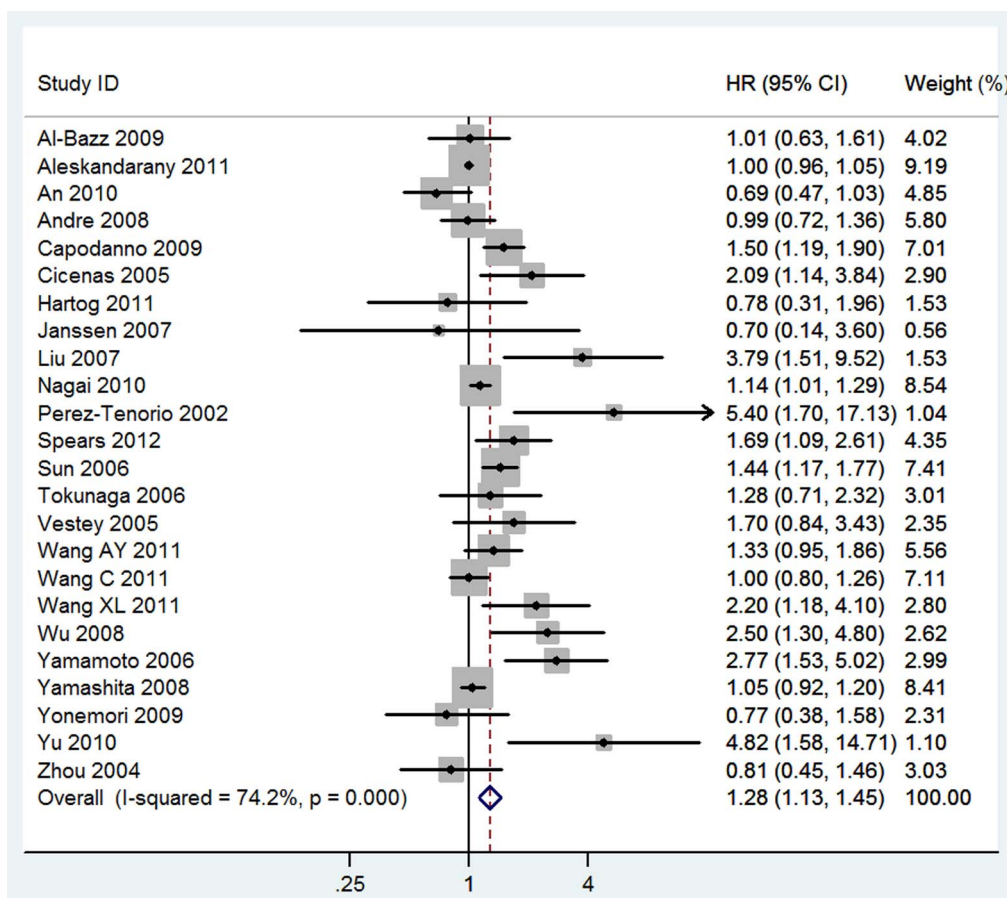
Specifically, we noted that the scoring methods used by published studies comprised of: (i) staining intensity alone; (ii) the proportion of tumor cells with positive staining alone; (iii) staining intensity score multiplied by or plus the score for the proportion of tumor cells with positive staining; or (iv) other derivative methods (for details, see Supplementary Table S2). With regard to each method, the cutoffs or thresholds used to define pAkt overexpression also varied. Due to the highly heterogeneous nature of these methods and definitions, we were unable to conduct meaningful subgroup or meta-regression analysis to investigate their impact on the observed prognostic effect of pAkt overexpression. This highlights

the need for a standardized methodology for pAkt status testing before this biomarker can be applied to clinical practice.

The funnel plots corresponding to Figure 2 and Figure 3 were asymmetrical, which indicated that our meta-analyses might have suffered from publication bias. However, there were alternative explanations for this, as studies have shown that the asymmetry of funnel plots could be due to other reasons than publication bias, such as true heterogeneity of effects, poor study quality, and the play of chance<sup>45</sup>. In view of the significant heterogeneity present in our meta-analyses, it is reasonable to say that publication bias possibly but not necessarily existed. Even if publication bias indeed occurred, our analysis showed that the summary HRs with publication bias adjusted for by trim-and-fill method were still statistically significant. Thus, we argue that publication bias did not constitute a major problem in the interpretation of our results.

Our finding about the adverse prognostic effect of pAkt overexpression is consistent with the observations in other solid tumors. For example, in the study of Nakanishi et al with 135 hepatocellular carcinoma patients, multivariate analysis identified pAkt overexpression as a strong predictor for early disease recurrence (relative risk: 12.5, 95% CI: 2.59–60.55) and poor prognosis (relative risk: 7.90, 95% CI: 1.25–50.00)<sup>9</sup>. The study of Cinti et al in 50 advanced gastric carcinomas showed that the five-year survival rate was 18% in the patients with pAkt overexpression versus 58% in the pAkt-negative ones<sup>46</sup>. These findings together with ours suggest that pAkt overexpression could be a common prognostic factor shared by multiple types of human cancer, and thus it has the potential for being a therapeutic target of great clinical significance.

In fact, preclinical studies have provided evidence that inhibiting Akt activation while giving other treatments might enhance the overall efficacy. For example, Chen et al showed that inhibition of



**Figure 3** | Meta-analysis of the association between pAkt overexpression and disease-free survival in breast cancer. Results are presented as individual and pooled HRs with corresponding 95% CIs. HR > 1 means that disease-free survival of the patients with pAkt overexpression is worse than that of the pAkt-negative ones, while HR < 1 means the opposite.

Akt activation by recombinant VP1 suppresses the progression of hepatocellular carcinoma<sup>47</sup>. Other Akt inhibitors such as RX-0201, PBI-05204 and GSK2141795 have also demonstrated activity in various solid tumors in preclinical and phase I studies<sup>48</sup>. In breast cancer, *in vitro* and *in vivo* studies have showed that Akt inhibitor MK-2206, alone or in combination with chemotherapy, has antitumor activity and may augment the efficacy of existing cancer therapeutics<sup>49,50</sup>. Currently, MK-2206 is undergoing phase II trials<sup>51</sup>. Results of these studies should shed new lights on the clinical utility of pAkt testing. If the drugs targeted at pAkt proved effective, pAkt expression status could be a potential predictive biomarker and thus used to make the treatment of breast cancer more individualized in the future, similar to the role of *EGFR* mutation status in the *EGFR*-targeted treatment of non-small cell lung cancer<sup>52</sup>.

In conclusion, this systematic review suggests that pAkt overexpression is an adverse prognostic factor in breast cancer in terms of both overall survival and disease-free survival. To facilitate its application, efforts are needed to develop a standardized assay methodology and to further evaluate the efficacy of Akt inhibition with regard to other treatments in clinical settings.

## Methods

**Literature search.** We performed a systematic search of PubMed, EMBASE (including the conference proceedings of American Society of Clinical Oncology and European Society of Medical Oncology) and Chinese Biomedical Literature Database (in Chinese) from their respective inception through 2013. The keywords used to search relevant publications included: “breast cancer\*”, “breast carcinoma\*”, “breast tumor\*”, “breast tumour\*”, “Akt\*”, “pAkt”, “p-Akt”, “prognos\*”, “outcome\*”, “progress\*”, “metasta\*”, “relapse\*”, “recurren\*”, “surviv\*”, “death\*”, “die\*”, “dead”, “dying”, “mortality”. As the association of pAkt status with prognosis was often investigated by secondary analysis in the studies that focused on PTEN protein

and/or *PIK3CA* gene, the following keywords related to the two biomarkers were also used in our literature search: “phosphatase and tensin homolog”, “PTEN”; “PIK3CA”, “PI3K\*”, “PIK3\*”, “phosphoinositide 3-kinase”, “phosphoinositide-3-kinase”, “phosphatidylinositol 3 kinase”, “phosphatidylinositol 3-kinase”, “PI 3-kinase”, “phosphatidylinositol-3 kinase”. No restrictions were placed on language or publication status. Wherever possible, the searches were limited to “human studies”. The reference lists of eligible studies and relevant reviews were also scrutinized for additional eligible studies.

**Study selection.** The titles and abstracts of all identified records were screened to judge their relevance. The full texts of the studies seemingly fulfilling the inclusion criteria were obtained for further assessment. Cohort studies that met all of the following criteria were considered eligible: (i) The subjects were patients diagnosed with breast cancer. (ii) The outcomes included overall survival, disease-free survival, or both. (iii) pAkt status was tested and correlated with the outcomes. Duplicates and studies with non-extractable data were excluded.

**Data extraction.** The following data were extracted from eligible studies: (i) bibliographic information, such as first author, country and publication year; (ii) data on clinical and pathological characteristics of patients, such as sample size, stage of disease, ER status, PR status, HER2 status, and treatments given; (iii) the proportion of patients with pAkt overexpression; (iv) main results of the study, such as HR and 95% CI (if available, multivariate estimates were preferable); (v) the information related to study quality (see below).

Authors of the original studies were contacted as needed to clarify the ambiguities in the reported methods or results and to seek additional data not included in the published reports. If not explicitly reported in the original paper and still not available after contact with author, HR was estimated according to the survival curves using the method developed by Parmar et al and recommended by the Cochrane Handbook for Systematic Reviews<sup>53</sup>. In rare cases, HR was not estimable and rate ratio was used as a substitute for it<sup>54</sup>. Data extraction was completed independently by two reviewers (Z.Y.Y. & M.Y.D.). Disagreements between the two were resolved by revisiting the original paper and discussion until consensus was reached.

**Quality assessment.** The quality of included studies was assessed according to the Newcastle-Ottawa scale<sup>55</sup>, which was frequently employed by previous studies<sup>56</sup>. This



Table 2 | Results of subgroup and meta-regression analyses

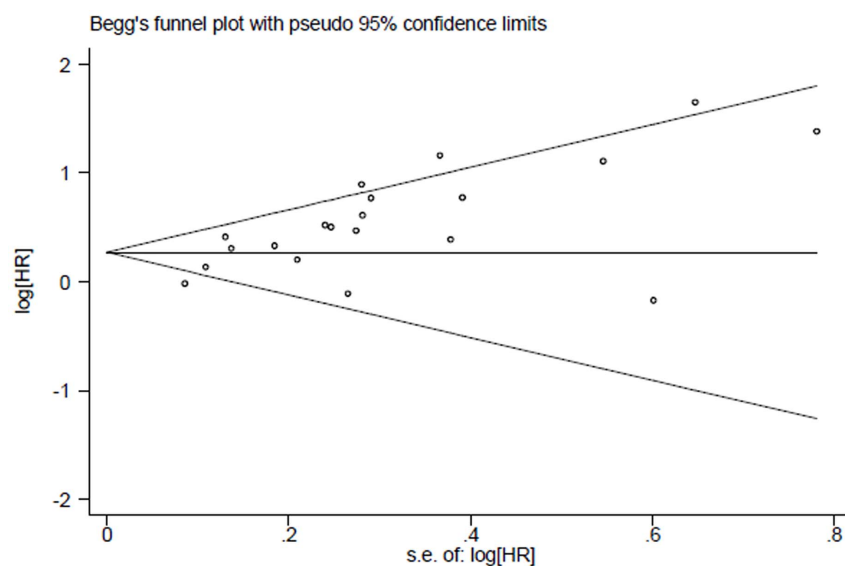
Outcomes, factors and subgroups	No. of studies	No. of patients	Summary HR (95% CI)	Heterogeneity	Meta-regression P-value
Overall survival:	20	6349	1.52 (1.29–1.78)	$I^2 = 58.4\%$ , $P = 0.001$	
1. Population					
Asian	9	2089	1.71 (1.30–2.26)	$I^2 = 77.4\%$ , $P = 0.000$	0.430
Non-Asian	11	4260	1.39 (1.20–1.61)	$I^2 = 0.0\%$ , $P = 0.515$	
2. Sample size					
<142	11	1078	1.72 (1.35–2.20)	$I^2 = 38.8\%$ , $P = 0.090$	0.183
$\geq 142$	9	5271	1.36 (1.13–1.64)	$I^2 = 60.3\%$ , $P = 0.010$	
3. ER-positive patients					
<50%	5	518	1.68 (1.16–2.44)	$I^2 = 45.9\%$ , $P = 0.116$	0.353
$\geq 50\%$	12	5370	1.37 (1.14–1.64)	$I^2 = 55.1\%$ , $P = 0.011$	
4. PR-positive patients					
<50%	6	1473	1.48 (1.02–2.15)	$I^2 = 53.4\%$ , $P = 0.057$	0.680
$\geq 50\%$	5	2495	1.61 (1.21–2.16)	$I^2 = 39.0\%$ , $P = 0.161$	
5. HER2-positive patients					
<50%	10	5198	1.50 (1.20–1.88)	$I^2 = 67.4\%$ , $P = 0.001$	0.557
$\geq 50\%$	5	486	1.74 (1.27–2.39)	$I^2 = 0.0\%$ , $P = 0.500$	
6. pAkt overexpression rate					
<50.4%	11	3741	1.69 (1.33–2.14)	$I^2 = 61.5\%$ , $P = 0.004$	0.690
$\geq 50.4\%$	8	2330	1.50 (1.27–1.77)	$I^2 = 0.0\%$ , $P = 0.793$	
7. Hormonal treatment					
Yes	1	130	2.17 (1.01–4.68)	NA	0.452
No	19	6219	1.50 (1.28–1.76)	$I^2 = 59.0\%$ , $P = 0.001$	
8. Trastuzumab					
Yes	3	261	1.63 (0.94–2.81)	$I^2 = 20.4\%$ , $P = 0.285$	0.801
No	17	6088	1.51 (1.21–1.78)	$I^2 = 62.1\%$ , $P = 0.000$	
9. Follow-up length					
<5 years	5	441	1.78 (1.22–2.61)	$I^2 = 22.5\%$ , $P = 0.271$	0.398
$\geq 5$ years	15	5908	1.47 (1.24–1.74)	$I^2 = 61.9\%$ , $P = 0.001$	
10. Effect measure					
Hazard ratio	18	6171	1.51 (1.27–1.81)	$I^2 = 60.1\%$ , $P = 0.001$	0.766
Risk ratio	2	178	1.56 (1.24–1.97)	$I^2 = 0.0\%$ , $P = 0.523$	
11. Analyzing method					
Univariate	14	2960	1.56 (1.26–1.93)	$I^2 = 66.7\%$ , $P = 0.000$	0.861
Multivariate	6	3389	1.40 (1.12–1.74)	$I^2 = 24.1\%$ , $P = 0.253$	
12. Study quality score					
<7	12	2594	1.37 (1.13–1.66)	$I^2 = 54.5\%$ , $P = 0.012$	0.159
$\geq 7$	8	3755	1.81 (1.36–2.42)	$I^2 = 61.0\%$ , $P = 0.012$	
Disease-free survival:	24	8683	1.28 (1.13–1.45)	$I^2 = 74.2\%$ , $P = 0.000$	
1. Population					
Asian	13	3272	1.36 (1.09–1.69)	$I^2 = 75.2\%$ , $P = 0.000$	0.796
Non-Asian	11	5411	1.24 (1.05–1.47)	$I^2 = 71.2\%$ , $P = 0.000$	
2. Sample size					
<142	11	1095	1.70 (1.28–2.26)	$I^2 = 58.5\%$ , $P = 0.007$	0.061
$\geq 142$	13	7588	1.13 (1.01–1.28)	$I^2 = 70.6\%$ , $P = 0.000$	
3. ER-positive patients					
<50%	2	154	1.32 (0.47–3.69)	$I^2 = 78.8\%$ , $P = 0.030$	0.796
$\geq 50\%$	19	7907	1.18 (1.05–1.33)	$I^2 = 68.5\%$ , $P = 0.000$	
4. PR-positive patients					
<50%	6	2086	1.07 (0.82–1.40)	$I^2 = 57.8\%$ , $P = 0.037$	0.315
$\geq 50\%$	8	4210	1.45 (1.03–2.04)	$I^2 = 74.2\%$ , $P = 0.000$	
5. HER2-positive patients					
<50%	16	7679	1.21 (1.05–1.40)	$I^2 = 73.2\%$ , $P = 0.000$	0.569
$\geq 50\%$	3	264	1.08 (0.60–1.95)	$I^2 = 26.8\%$ , $P = 0.255$	
6. pAkt overexpression rate					
<50.4%	13	4438	1.38 (1.14–1.67)	$I^2 = 66.6\%$ , $P = 0.000$	0.575
$\geq 50.4\%$	10	3967	1.27 (0.98–1.63)	$I^2 = 77.0\%$ , $P = 0.000$	
7. Hormonal treatment					
Yes	3	295	2.74 (1.16–6.52)	$I^2 = 74.4\%$ , $P = 0.020$	0.067
No	21	8388	1.21 (1.07–1.36)	$I^2 = 69.4\%$ , $P = 0.000$	
8. Trastuzumab					
Yes	1	44	0.77 (0.38–1.58)	NA	0.322
No	23	8639	1.30 (1.14–1.47)	$I^2 = 75.1\%$ , $P = 0.000$	
9. Follow-up length					
<5 years	6	1479	1.60 (0.88–2.89)	$I^2 = 78.5\%$ , $P = 0.000$	0.572
$\geq 5$ years	18	7204	1.24 (1.10–1.41)	$I^2 = 73.4\%$ , $P = 0.000$	
10. Effect measure					
Hazard ratio	18	5998	1.42 (1.14–1.75)	$I^2 = 71.2\%$ , $P = 0.000$	0.435
Risk ratio	6	2685	1.21 (1.03–1.42)	$I^2 = 71.3\%$ , $P = 0.000$	



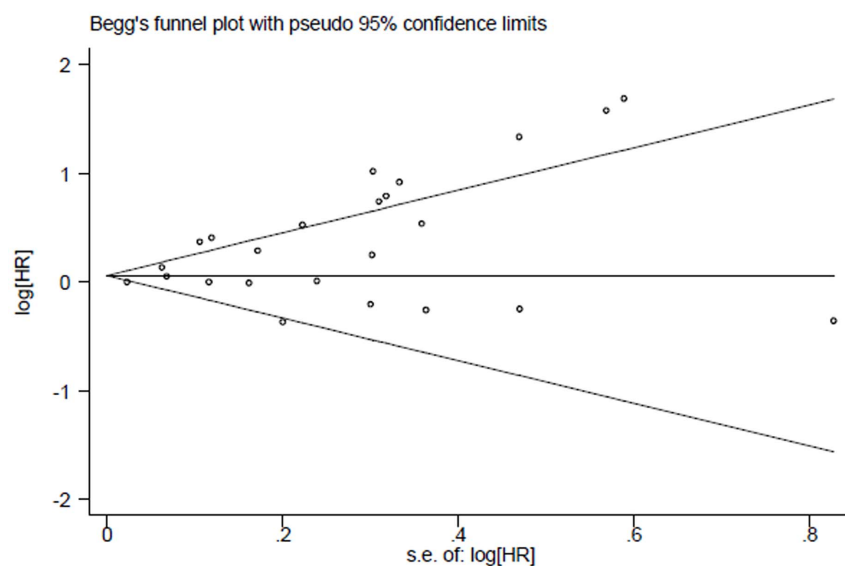
Table 2 | Continued

Outcomes, factors and subgroups	No. of studies	No. of patients	Summary HR (95% CI)	Heterogeneity	Meta-regression P-value
11. Analyzing method					
Univariate	12	3886	1.16 (1.02–1.31)	$I^2 = 68.8\%$ , $P = 0.000$	0.106
Multivariate	12	4797	1.65 (1.22–2.22)	$I^2 = 76.1\%$ , $P = 0.000$	
12. Study quality score					
<7	9	2397	1.14 (1.00–1.31)	$I^2 = 33.1\%$ , $P = 0.153$	0.403
$\geq 7$	15	6286	1.41 (1.15–1.71)	$I^2 = 81.4\%$ , $P = 0.000$	

Abbreviations: HR = hazard ratio; CI = confidence interval; ER = estrogen receptor; PR = progesterone receptor; HER2 = human epi-dermal growth factor receptor 2; NA = not applicable.



(A)



(B)

**Figure 4** | Funnel plots to examine the possibility of publication bias in the data for overall survival (A) and that for disease-free survival (B). The standard error of log HR (S.E. of log HR) was plotted against log HR for each individual study as represented in a circle. Egger's tests showed that the funnel plots were asymmetric ( $P < 0.001$  for Figure 4(A);  $P = 0.002$  for Figure 4(B)), which could be due to potential publication bias among other reasons.



scale focuses on three aspects of studies, including selection of patients, comparability of baseline characteristics, and outcome assessment. For each aspect, there are 1–4 items for detailed evaluation. The study quality was denoted by a numerical score ranging from 0 to 9, with 9 representing the highest quality. Quality assessment was completed independently by two reviewers (J.Q.Y. & C.M.). Disagreements between the two were resolved by revisiting the original paper and discussion. Unsettled disagreements were referred to a third researcher for final decision (J.L.T.).

**Statistical analysis.** The primary and secondary clinical outcomes of our interest were overall survival and disease-free survival, respectively. The effect of pAkt overexpression on the outcomes was measured by HR with 95% CI, and the HRs from relevant studies were combined to produce a summary HR for each outcome. HR > 1 means that the prognosis of patients with overexpressed pAkt is worse than that of the other patients, while HR < 1 means the opposite. The statistical heterogeneity among studies was assessed by the Cochran's Q test and the I<sup>2</sup> statistic<sup>57,58</sup>. A p value ≤ 0.10 for the Q test or an I<sup>2</sup> > 50% was suggestive of substantial between-study heterogeneity. In case of substantial heterogeneity, the random-effects model was used for meta-analysis; otherwise, the fixed-effects model was used. If substantial, the heterogeneity was investigated by subgroup and meta-regression analyses to see if it could be explained by the following factors: population, sample size, the respective proportion of subjects positive for ER, PR and HER2, pAkt overexpression rate, hormonal treatment, trastuzumab treatment, length of follow-up, effect measure used in original studies (HR vs. rate ratio), analyzing method (univariate vs. multivariate) and study quality. Sensitivity analyses were conducted by omitting one study each time and by using the alternative analysis model (e.g. switching from the random-effects model to the fixed-effects model). Begg's funnel plot and Egger's test were used to examine the possibility of publication bias if a meta-analysis included 10 or more studies<sup>59</sup>. In presence of an asymmetric funnel plot, the Duval and Tweedie nonparametric trim-and-fill method was used to adjust for the potential publication bias and to obtain an adjusted summary HR from the corresponding meta-analysis<sup>60</sup>. All the analyses were performed with STATA software, version 11.0 (StataCorp, College Station, TX, USA).

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## Additional information

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