High-Confidence Off-Policy Evaluation

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AAAI-15. 29th AAAI Conference on Artificial Intelligence, Austin, Texas, USA, January 2015.

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Overview

This paper develops a method to compute the confidence that the expected return of a policy exceeds some lower bound, using only trajectories generated from other policies.

Terminology:

- Evaluation Policy The policy which we wish to estimate the expected return of.
- Behavior Policy The policies used to estimate the return of the evaluation policy.

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Motivation

"..execution of a new policy can be costly or dangerous if it performs worse than the policy that is currently being used..."

Examples:

- News recommendation systems
- Patient diagnosis systems
- Neuroprosthetic control
- Automatic drug administration

Problem Setup

This work follows the standard Markov Decision Process (MDP) formalism:

- S: State space.
- A: Action space.
- $r_t \in [r_{max}, r_{min}]$: Reward at time t.
- $\gamma \in [0, 1]$: Discount factor.
- π(a | s, θ): Probability of taking action a in state s given policy parameters θ.

• $\tau = \{s_1, a_1, r_1, \dots, s_T, a_T, r_T\}$: A trajectory.

Problem Setup

Define the normalized and discounted return of a trajectory to be:

$$R(\tau) = \frac{(\sum_{t} \gamma^{t-1} r_{t}) - R_{-}}{R_{+} - R_{-}}$$

Where R_+ and R_- are upper and lower bounds for $\sum_t \gamma^{t-1} r_t$.

Ideally, we want to know the expected return given the *evaluation* policy parameters θ :

$$\rho(\theta) = \mathbb{E}\left[R(\tau) \,|\, \theta
ight]$$

Generating Unbiased Estimates of $\rho(\theta)$

Key Idea

Given a dataset $\mathcal{D} = \{(\tau_i, \theta_i) : \tau_i \text{ generated using } \theta_i\}$ estimate $\rho(\theta)$ using *importance sampling*.

Importance Sampling

Given a *target distribution* p and *sampling distribution* q and a function f:

$$\mathbb{E}_{x\sim p}(f(x)) = \frac{1}{n} \sum_{i=1}^{n} f(x_i) \frac{p(x_i)}{q(x_i)}$$

where $x_i \sim q$.

Generating Unbiased Estimates of $\rho(\theta)$

In the context of our problem, the *target distribution* is $Pr(\tau_i | \theta)$ (e.g. the probability of a trajectory under the *evaluation policy*), and the *sampling distribution* is $Pr(\tau_i | \theta_i)$ (e.g. the probability of a trajectory under the *behavior policy* it was generated by).

The expected return of the *evaluation policy* can then be estimated by:

$$\rho(\theta) \approx \frac{1}{n} \sum_{i=1}^{n} R(\tau_i) \frac{\Pr(\tau_i \mid \theta)}{\Pr(\tau_i \mid \theta_i)}$$
$$= \frac{1}{n} \sum_{i=1}^{n} \underbrace{R(\tau_i) \prod_t \frac{\pi(a_t \mid s_t, \theta)}{\pi(a_t \mid s_t, \theta_i)}}_{\hat{\rho}(\theta, \tau_i, \theta_i)}$$

Interlude

Recall our motivation:

"..execution of a new policy can be costly or dangerous if it performs worse than the policy that is currently being used..."

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In the following slides, we will review a few "classical" results providing lower bounds for the expectation of a random variable that can be estimated using samples.

We introduce the following variables/assumptions:

- X_i: real-valued, positive, bounded random variables (e.g. importance weighted returns).
- μ : $\mathbb{E}(X_i)$ for all X_i .
- b: A real-number satisfying $Pr[X_i < b] = 1$ for all X_i .

"Classical" Results

Chernoff-Hoeffding (CH) inequality

With probability at least $1 - \delta$:

$$\mu \geq \frac{1}{n} \sum_{i=1}^{n} X_i - b \sqrt{\frac{\ln\left(1/\delta\right)}{2n}}$$

Anderson (AM) inequality

With probability at least $1 - \delta$:

$$\mu \ge z_n - \sum_{i=0}^{n-1} (z_{i+1} - z_i) \min\left\{1, \frac{i}{n} + \sqrt{\frac{\ln(2/\delta)}{2n}}\right\}$$

where z_i are the samples X_i in increasing order, and $z_0 = 0$.

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Maurer & Pontil's empirical Bernstein (MPeB) inequality With probability at least $1 - \delta$:

$$\mu \geq \frac{1}{n} \sum_{i=1}^{n} X_i - \frac{7b \ln (2/\delta)}{3(n-1)} - \frac{1}{n} \sqrt{\frac{\ln (2/\delta)}{n-1} \sum_{i,j=1}^{n} (X_i - X_j)^2}$$

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New Result

The author's use the **MPeB** inequality to prove the following: Theorem 1 Let $Y_i = \min\{X_i, c_i\}$ where $c_i > 0$, with probability at least $1 - \delta$:

$$\mu \ge \left(\sum_{i=1}^{n} \frac{1}{c_i}\right)^{-1} \sum_{i=1}^{n} \frac{Y_i}{c_i} - \left(\sum_{i=1}^{n} \frac{1}{c_i}\right)^{-1} \frac{7n \ln (2/\delta)}{3(n-1)} \\ - \left(\sum_{i=1}^{n} \frac{1}{c_i}\right)^{-1} \sqrt{\frac{\ln (2/\delta)}{n-1} \sum_{i,j=1}^{n} \left(\frac{Y_i}{c_i} - \frac{Y_j}{c_j}\right)^2}$$

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- ► These inequalities can be inverted so that given a lower bound μ_{-} we can determine the confidence (e.g. value of δ) that $\mu \ge \mu_{-}$.
- Theorem 1 requires pre-specified thresholds c_i. In the paper, the authors select a single value c* and set all c_i = c*. They show how an optimal value of c* can be determined by splitting the dataset and performing cross-validation.

Demo https://www.youtube.com/watch?v=x_qDs2kA7H4&feature= youtu.be

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Method	ho
Thm. 1	0.154
CH	-5,831,000
MPeB	-129,703
AM	0.055

Table: 95% confidence lower bounds

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Experiment 2: Digital Marketing using Real-World Data

Problem

Decide the optimal policy for user-specific targeting of advertisements. Rewards are measured by advertisement click rates (e.g. agent recieves +1 if a user clicks the ad, and 0 otherwise).

Data

- From a website for a Fortune 50 company.
- \blacktriangleright > 100,000 visitors a day.
- Each user has 31 features.
- Agent must select from two clusters of advertisements.

Experiment 2: Digital Marketing using Real-World Data



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Experiment 2: Digital Marketing using Real-World Data



Black - 2 million off-policy trajectories Blue - 5 million off-policy trajectories Red - 5 million off-policy trajectories + 1 million on-policy trajectories

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